# Partners Risk Perception and Couple Portfolio Allocation* 

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#### Abstract

This work investigates the relationship between households' portfolio allocation and partners' risk preferences using the English Longitudinal Study of Ageing data. I develop a portfolio choice model where partners pooled their risk preferences and then chose the optimal portfolio allocation as a single decision unit. Then, I study the determinants of household portfolio allocation controlling for selection into stock market participation. Last, I compare the goodness of fit of the proposed pooled approach and the standard unitary model, showing that the former fits significantly better the data.


Keywords household portfolio; risk aversion; group choice.

[^0]
## 1 Introduction

The workflow of the household decision-making process is of core interest in economics, and understanding its mechanisms might spread new light on consumers behaviours. In this context, financial choices are of particular interest because of their potential impacts on households present and future economic status. Moreover, the responsibility of these choices largely falls into individuals' hands, thus, they are good candidates to study how individual preferences affect household outcomes. Together with housing, portfolio allocation is the most important financial choice of households. It was broadly studied over the past years, and the models used predict each household holds a fraction of its wealth in risky assets if the equity premium is positive (e.g.: Samuelson (1975) and Merton (1969)). These models rely on the so-called unitary approach, which considers the household as a unique decision unit that behaves as a single agent with well-defined preferences.
Chiappori (1988) shades new light on the household (and group) decision-making process, introducing the collective model. In this approach, households behave as a multi-dimension system of several members, which may show different preferences. An intrahousehold bargaining process is assumed to take place among the household members, which combines individual preferences and drives the final choice. The collective model is largely used in studies concerning household labour supply (e.g.: Chiappori (1988), Chiappori (1992)), consumption choices (e.g.: Cherchye et al. (2007), van Leeuwen et al. (2020)) and household production decisions (Apps and Rees (1997)), but to date only a few studies household portfolio allocation under this perspective (see Gomes et al. (2020)). Addoum et al. (2016) study the connection between marital decisions, consumption, and household investments. They show that changes in marital status or spouses' relative income imply a significant reallocation of the household portfolio. Olafsson and Thornquist (2018) use the potential earnings of spouses, instead of actual earnings, as a proxy of the household decision-makers bargaining powers. In line with Addoum et al. (2016), they show that the higher the weight of the wife, the lower the household probability of holding equity. In other words, if the female partner has higher decision power, the household portfolio is less risky. Last, Gu et al. (2021) investigate the gender gap of bargaining power in the household portfolio decision-making process. Their results show that the household investments reflect the preferences of the male partner $44 \%$ more than the female partner characteristics, with gender norms that play a relevant role in explaining this difference.

This paper assumes that the two partners bargain about risk preferences, which play a crucial role in portfolio choices as largely documented in the literature (e.g.: Zhang (2017)). The theoretical model the bargaining process only concerns household risk preferences. The outcome of this process drives the household portfolio choice, where households decide about the convenience of stock market participation and optimal wealth allocation.
I empirically test the implication of the model using the English Longitudinal Study of Ageing (ELSA) panel dataset. The estimates are in line both with the implication of the model and the results of the literature (see Guiso et al. (2003) for stock market participation and Jappelli and Padula (2015) for risky assets share). Then, I compare the goodness of fit of the standard unitary approach with the proposed collective approach. The results show that the collective model fits significantly better, thus, it is relevant to account for the preferences of all the household decision-makers when studying portfolio allocation.
This paper contributes to the household portfolio choice models with limited stock market participation (Gomes and Michaelides's (2005); Wachter and Yogo (2010)). Generally, households
are considered single decision-making units, while large stock market non-participation rates are designed as a consequence of stock market participation costs, lack of financial literacy, and so on. The literature refers to these models as unitary models, where the preferences of the decision-making unit are approximated with those of the male partner or the household head. Vissing-Jorgensen (2002) introduces three types of participation costs: fixed or lump-sum entry costs, variable transaction costs and per-period trading costs.
I introduce a portfolio-choice model that accounts for household risk preferences measured as a weighted average of partners' risk tolerance. The weights can be interpreted as each agent's bargaining power. Therefore, the partner who holds the "purse strings" (Bertocchi et al. (2014)) would have a higher decision power and influence in the determination of the household risk tolerance.

The rest of the paper is organized as follow: Section 2 introduces the collective household model for household's portfolio allocation. Section 3 describes the ELSA dataset, Section 4 presents the empirical analysis and Section 5 concludes.

## 2 Theoretical model

This section introduces the theoretical model that studies the heterogeneity of household portfolio choice considering the risk preferences of both household decision-makers. The model assumes that the household members first bargain on risk preferences, and then the results of this bargain drives the portfolio allocation choice. I use a mean-variance utility function ${ }^{3}$ where the household risk tolerance is measured as a weighted average of partners risk preferences.

### 2.1 Collective model: household utility and weighted risk aversion

Assume that the economy has only two assets: a risk-free asset (representative of the treasury bill market) and a risky asset (representative of the stock market). The two assets have different expected returns: the risk-free asset has certain returns $r$ while the risky asset has returns $\tilde{r}=r+\tilde{s}$, where $\tilde{s} \sim N\left(\mu_{s}, \sigma_{s}^{2}\right)$. Each household $h$ has two decision makers (partners), agents $a$ and $b$, and holds the initial wealth $w_{h}$. The household wants to maximize its utility $u\left(W_{h}\right)$, where $W_{h}$ represents the expected household wealth after assets returns. The crucial decision is about $\alpha$, that identifies the proportion of wealth $w_{h}$ allocated in risky assets.
The household wealth $W_{h}$ is:

$$
\begin{equation*}
W_{h}=w_{h}[(1-\alpha)(1+r)+\alpha(1+\tilde{r})]=w_{h}[(1+r)+\alpha \tilde{s}] \tag{1}
\end{equation*}
$$

Thus, the household utility maximization program can be written in terms of the value function $V_{h}$ as a function of $\alpha$ :

$$
\begin{equation*}
\max _{\alpha} V_{h}(\alpha)=\max _{\alpha} u\left(W_{h}\right)=\max _{\alpha} u\left(w_{h}[(1+r)+\alpha \tilde{s}]\right) \tag{2}
\end{equation*}
$$

Assuming CARA (exponential) utility function of the form $u(z)=-e^{-\rho z}$, where $\rho$ is the absolute risk aversion coefficient, household utility becomes:

$$
\begin{equation*}
u\left(W_{h}\right)=-e^{-\rho_{h} W_{h}} \tag{3}
\end{equation*}
$$

[^1]where $\rho_{h}$ identifies the household risk preference. $\rho_{h}$ is a weighted sum of the partners risk preferences, where the weights represents the bargaining power of each partner in the decision making process. Therefore, household risk aversion is:
\[

$$
\begin{equation*}
\rho_{h}=\mu_{a} \rho_{a}+\mu_{b} \rho_{b} \tag{4}
\end{equation*}
$$

\]

where $\rho_{a, b}$ are the risk preferences of the household decision makers, and $\mu_{a, b}$ are the bargaining powers. I normalize the weights as follow:

$$
\begin{equation*}
\gamma=\frac{\mu_{a}}{\mu_{a}+\mu_{b}} \quad(1-\gamma)=\frac{\mu_{b}}{\mu_{a}+\mu_{b}} \tag{5}
\end{equation*}
$$

such that $\gamma \in[0,1]$ and $\rho_{h}$ becomes:

$$
\begin{equation*}
\rho_{h}=\gamma \rho_{a}+(1-\gamma) \rho_{b} \tag{6}
\end{equation*}
$$

Thus, $V_{h}(\alpha)$ in Equation 2 becomes:

$$
\begin{align*}
\max _{\alpha} V_{h}(\alpha) & =\max _{\alpha} E\left[u\left(W_{h}\right)\right]=\max _{\alpha} E\left[-e^{-\rho_{h} W_{h}}\right] \\
& =\max _{\alpha} \rho_{h} w_{h}\left[(1+r)+\alpha \mu_{s}-\frac{1}{2} \alpha^{2} \sigma_{s}^{2} \rho_{h} w_{h}\right] \tag{7}
\end{align*}
$$

Solving the first order condition for $\alpha$, the optimal share of wealth allocated in risky assets is:

$$
\begin{equation*}
\alpha=\frac{\mu_{s}}{\sigma_{s}^{2} w_{h} \rho_{h}} \tag{8}
\end{equation*}
$$

The optimal $\alpha$ is proportional to the risk premium $\mu_{s}$ and decreasing in the variance (risk) of returns and in household risk aversion.
Under the condition of Equation 8 and with the strong assumption that all households have the same information about stock market returns, the heterogeneity in households' portfolio depends on wealth and risk preferences ${ }^{4}$.

### 2.2 Preference shifter

The solution proposed in Equation 8 implies that the heterogeneity in $\alpha$ depends entirely on risk preferences and household wealth, assuming that individuals have common priors (i.e., they all experience the same stock expected returns and return variance). However, it is unlikely that household demographics such as education, income or age do not affect household portfolio decision.
I allow household characteristics to affect the portfolio choice process through risk preferences. Thus, risk preferences now are a function of household and partners demographics, $\rho_{i}=\rho(z)$, where $i=a, b$ and $z=\beta x_{i}+\phi h+\epsilon . x$ is a vector of individual demographics (e.g.: ages, education), $h$ is a vector of household demographics (e.g.: aggregate income, children) and $\beta$ and $p h i$ are parameter matrices and $\epsilon$ is the i.i.d. error term that represents the unobserved variation in taste shifts across households. Therefore, the maximization problem in Equation 7 becomes:

$$
\begin{equation*}
\max _{\alpha} V_{h}(\alpha)=\max _{\alpha}\left[\rho_{h}(z)\right] w_{h}\left[(1+r)+\alpha \mu_{s}-\left(\frac{1}{2} \alpha^{2} \sigma_{s}^{2} w_{h}\left[\rho_{h}(z)\right]\right)\right] \tag{9}
\end{equation*}
$$

[^2]and the optimal $\alpha$ is:
\[

$$
\begin{equation*}
\alpha=\frac{\mu_{s}}{\sigma_{s}^{2} w_{h}\left[\rho_{h}(z)\right]} \tag{10}
\end{equation*}
$$

\]

Equation 8 and Equation 10 show that the optimal proportion of wealth allocated in risky assets depends on risk preference, stock market returns and wealth itself. However, the former implies that household or individual demographics do not affect portfolio choices, while a large amount of literature shows that they influence household finances. Education (Cooper and Zhu (2016), Poterba et al. (2013)), health (Poterba et al. (2013)), age (Ameriks and Zeldes (2004), Bertocchi et al. (2014)), wealth (Wachter and Yogo (2010)) and financial literacy (Jappelli and Padula (2015), Lusardi (2008)) are only some of the determinants of household portfolio decision. Thus, non including them in the model may produce severe bias in the estimation.
The solution proposed in Equation 10 shows that the optimal share of wealth allocated in risky assets is positive for each household. In other words, if the risk premium $\mu_{s}$ is positive, every households invests a fraction of its financial wealth in the risky asset, as in Samuelson (1975) and Merton (1969). However, a large fraction of households does not hold stocks: this is the so-called stock holding puzzle (e.g.: Guiso et al. (2003) describe and discuss this issue across European countries).

### 2.3 Introducing stock market participation costs

One of the possible explanations of the stock holding puzzle are stock market participation costs. They reduce risky assets' expected returns and increase the probability of losses. Therefore, depending on the amount that they have to pay, the households may decide to not participate. VissingJorgensen (2002) identifies three main types of costs that affect the stock market participation choice: fixed entry costs (learning about financial markets), variable transaction costs (trading fees or bid-ask spread) and per period trading costs (broker subscription or bank fees). Working with US data, Vissing-Jorgensen (2002) estimates that a relatively low per period cost (50\$ per month) explains the non participation of half of non-stockholders.

Including costs in the portfolio choice process means that the household utility maximization problem has now two steps: the first concerns the stock market participation decision (i.e.: $\alpha>0$ or $\alpha=0$ ) and the second solves the utility maximization problem, following Equation 10 . In case of costs, the expected household wealth after assets returns $W_{h}$ is described by the following framework:

$$
\begin{align*}
& \alpha=0 \rightarrow W_{h}=W_{h s}=w_{h}[(1+r)] \\
& \alpha>0 \rightarrow W_{h}=W_{h r}=w_{h}[(1-\alpha)(1+r)+\alpha(1+r+\tilde{s})]-C \tag{11}
\end{align*}
$$

where $C$ are the stock market fixed entry costs payed at the end of the period. The household evaluates whether holding risky assets is convenient or not, considering that in case of stock market participation it has to pay the lump sum cost $C . W_{h s}$ and $W_{h r}$ represent the expected household wealth at the end of the period in the two cases, non- and stock-holding, respectively. Then, the optimal amount allocated in risky assets $\alpha$ solves:

$$
\begin{equation*}
E\left[u^{\prime}\left(w_{h}((1+r)+\alpha \tilde{s})-C\right) \cdot w_{h}(\tilde{s})\right]=0 \tag{12}
\end{equation*}
$$

and $\alpha=0$ is a solution of the maximization problem if and only if the equity premium is 0 . I define the certainty equivalent of the risk premium $\tilde{s}$ as follow:

$$
\begin{equation*}
E\left[u\left(w_{h}((1+r)+\alpha \tilde{s})-C\right)\right]=u\left[w_{h}((1+r)+\alpha \hat{s})-C\right] \tag{13}
\end{equation*}
$$

where $\hat{s}$ represents the risk adjusted equity premium. Therefore, the household evaluates:

$$
\begin{align*}
& \mathbf{E}\left[u\left(W_{h r}\right)\right]>u\left(W_{h s}\right) \rightarrow \mathbf{E}\left[u\left(w_{h}((1+r)+\alpha \tilde{s})-C\right)\right]>u\left(w_{h}(1+r)\right) \\
& \rightarrow w_{h}(1+r)+w_{h} \alpha \hat{s}-C>w_{h}(1+r)  \tag{14}\\
& \rightarrow w_{h}>\frac{C}{\alpha \hat{s}}
\end{align*}
$$

The condition derived in Equation 14 defines a threshold of minimum wealth for potential investors that is proportional to fixed costs $C$, optimal share of wealth allocated in risky assets, $\alpha$, and the risk adjusted equity premium $\hat{s}$ :

$$
\begin{equation*}
\bar{w}=\frac{C}{\alpha \hat{s}} \tag{15}
\end{equation*}
$$

Then, when the household initial endowment of wealth is lower than $\bar{w}$, non-participation is the optimal choice, otherwise the household invests $\alpha$ share of its financial wealth in risky asset and pays $C$ at the end of the period.
Figure 2.1 simulates the wealth thresholds in equation 15 for different level of lump sum stock market entry costs and absolute risk aversion. The results assume constant household wealth, risk adjusted equity premium, expected returns and variance of the risky asset. Each color corresponds to a specific level of household risk aversion, thus, each dot corresponds to a different combination of costs and household risk. If the household corresponding dot is above the black line, the best choice is to not participate in the risky asset market. The figure shows how even very low participation costs ( $300 £$ at most paid once in a lifetime) imply low participation rate for mid-high level of risk aversion.

## 3 ELSA data

ELSA $^{5}$ is a longitudinal survey that collects data from a representative sample of English people aged $50+$. It is a biennial survey (first wave in 2002) that aims to gather data to study all the problems and aspects of ageing, like social care, retirement, pension policies and social participation. The original sample of ELSA (first wave) was selected from the Health Survey for England (HSE ${ }^{6}$ ) respondents in the period 1998-2001. After the first survey in 2002, younger age groups are refreshed to balance the panel over time.

This paper works with Wave 8 of ELSA, which collects data about 8445 individuals, interviewed between May 2016 and June 2017. Researchers introduced a series of new and innovative measures

[^3]Figure 2.1: Household stock market participation wealth threshold ( $\bar{w}_{h}=$ median hh financial wealth) by lump sum participation costs. Each color is associated with a specific level of risk aversion, and each dot corresponds to a combination of risk aversion and lump sum participation costs. When the dot is above the black line, the household does not participate in the stock market.

that have broadened the scope of the study. Among the new questions, Wave 8 includes three self-assessed measures of risk preferences: one related to the general propensity to take risks, one to financial risk taking and one to the health domain. The purpose of this paper is to study the household portfolio decision process and these questions are of particular interest because of the crucial role of risk tolerance in financial choices. Several papers point out that these qualitative questions predict behaviour across various domains (Caliendo et al. (2009), Fouarge et al. (2014)), including risk preferences, when experimental data (like lottery choices) are not available.
In ELSA, the participants answer to the following general risk tolerance question:
Are you generally a person who is fully prepared to take risk, or do you try to avoid taking risks?

The respondent chooses an integer between 0 (Avoid taking risks) and 10 (Fully prepared to take risks $]^{7}$ The predefined structure of the answers implies that they return a self-assessed measure of risk tolerance, rather than risk aversion. Then, these measures should be positively correlated with $\alpha$ according to Equation 10, where the risk measure, $\rho_{h}$, represents relative risk aversion.
The survey provides a second question, related to respondents patience:
Are you generally an impatient person, or someone who always shows great patience?

[^4]The respondent chooses an integer between 0 (Very impatient) and 10 (Very patient) ${ }^{8}$.
The correlation between partners' risk preferences (both general and financial) is always positive and significant, but relatively weak (coefficient vary between $7.7 \%$ and $17.7 \%$, depending on the specific item considered). The highest correlation is the one between the financial risk tolerance of husband and wife, that may be a signal of assortative matching of partners.

### 3.1 Sample selection and description

The survey distinguishes between three financial unit categories: singles, couples with separate finances and couples with joint finances. I use the data about individuals who are in a couple with joint finances aged less than 90 years. I select the male-female couples that have positive income (labour and pension income, including state benefit transfers), non-negative net financial wealth (the sum of savings and investments, subtracting financial debts from credit cards, overdrafts and other private debts but not mortgages) and share of net financial wealth allocated in risky assets between 0 and 1 (computed as the ratio between the amount of risky assets and the household net financial wealth). The final sample is composed of 1406 male-female couples, i.e. 2812 individuals ${ }^{9}$.
Table 1 presents the households' basic demographics, while Tables 2 and 3 show the sample statistics of partners. Table 2 distinguishes partners by gender, while Table 3 distinguishes partners by financial respondent (and non-financial respondent). The financial respondent is the partner that answers to the Income \& Asset section of the ELSA survey. Note that the selected households have joint finances: in these cases, the ELSA interviewer asks financial information only to one of the two spouses (financial respondent) and her/his answers are copied in the survey of the partner (non-financial respondent).

Table 1: ELSA Wave 8 summary statistics: couples with joint finances.

|  | All | Non-stockholders | Stockholders |
| :--- | :---: | :---: | :---: |
| hh obs | 1406 | 358 | 1048 |
| financial respondent: male | $60,9 \%$ | $55,0 \%$ | $62,9 \%$ |
| hh income: mean (weekly $£$ ) | 651,9 | 532,6 | 692,6 |
| hh income: median | 576,0 | 459,7 | 616,2 |
| hh income: std | 411,1 | 325,9 | 429,0 |
| net financial wealth: mean (thousand $£)$ | 153,6 | 45,4 | 190,5 |
| net financial wealth: median (thousand $£)$ | 66,7 | 13,6 | 101,3 |
| net financial wealth: std (thousand $£)$ | 267,9 | 182,9 | 282,0 |
| gross financial wealth: mean (thousand $£)$ | 154,5 | 46,0 | 191,5 |
| gross financial wealth: median (thousand $£)$ | 67 | 15 | 103 |
| gross financial wealth: std (thousand $£)$ | 267,8 | 182,9 | 281,8 |
| stock share of financial wealth | $32,8 \%$ | - | $44,0 \%$ |
| share of hh income : male $(\gamma)$ | $66,6 \%$ | $66,2 \%$ | $66,8 \%$ |
| share of hh income : respondent $(\gamma)$ | $56,2 \%$ | $54,7 \%$ | $56,7 \%$ |

Table 1 shows that around $75 \%$ of the households hold risky assets, where risky assets are defined as shares, bonds, stocks and shares ISAs or life insurance ISAs ${ }^{10}$. The share of stockholders is relatively high; however, ISAs products are widely spread among the English population and, included in

[^5]the definition of risky assets, consistently increase the participation rate of households. Labour and pension income are higher among stockholders, as well as net and gross household financial wealth. This is in line with the fixed entry costs assumption of Section 2.3. a higher household wealth corresponds to a lower impact of fixed entry costs on the portfolio returns, increasing the probability of participation ${ }^{11}$. The last two rows of Table 1 show the share of household labour and pension income of the financial respondent and male partner. On average, males contribute more to household income, with no difference between stockholders and non-stockholders. The financial respondent also generally holds a larger share of the household income, but the difference between the two partners is now ten percentage points lower. This is not a surprise, indeed around $40 \%$ of the financial respondent are females, which generally earn lower salary/pension.
Table 2 shows partner characteristics by gender. Comparing participants and non-participants households, a relevant difference concerns the education of partners: highly educated partners (college or higher education completed) are almost one-third among stockholders, while only one over ten non-stockholders complete college/university. With respect to risk preferences, males have higher risk tolerance than females. Wives show lower risk tolerance (general and financial) both when comparing stockholders and non-stockholders partners, in line with the literature. Moreover, the stockholders show a higher risk tolerance than non-stockholders, especially for financial risk preferences. Table 3 shows partner characteristics by financial and non-financial respondent. First, female financial-respondent are more among the non-stockholders, while the differences in age are negligible and the proportion of highly educated and low educated individuals is in line with Table 2. Concerning risk preferences, the differences between financial and non-financial respondents are tight: As mentioned above, this attenuation derives from the fact that $40 \%$ of the financial respondents are females, which shows a lower risk tolerance.

Table 2: ELSA Wave 8 summary statistics - male and female partners.

|  | All | Non-stockholders | Stockholders |
| :--- | :---: | :---: | :---: |
| age: male | 68,5 | 68,6 | 68,5 |
| age: female | 66,1 | 66,3 | 66,1 |
| low edu: male | $32,3 \%$ | $50,0 \%$ | $26,2 \%$ |
| mid edu: male | $42,0 \%$ | $38,3 \%$ | $43,2 \%$ |
| high edu: male | $25,7 \%$ | $11,7 \%$ | $30,5 \%$ |
| low edu: female | $29,8 \%$ | $43,9 \%$ | $25,0 \%$ |
| mid edu: female | $48,5 \%$ | $46,4 \%$ | $49,2 \%$ |
| high edu: female | $21,7 \%$ | $9,8 \%$ | $25,8 \%$ |
| general risk: male | 5,0 | 4,9 | 5,1 |
| general risk: female | 4,6 | 4,1 | 4,3 |
| financial risk: male | 3,6 | 3,2 | 3,7 |
| financial risk: female | 2,9 | 2,8 | 2,9 |

Table 4 gives an overview of the households wealth (and portfolio) allocation. Total household wealth is increasing in household income and is higher among stockholders. Comparing the first three quartiles of the income distribution, the wealth of stockholders is almost twice the wealth of non-stockholders. Last, the households with higher income are generally younger. On average, housing accounts for more than $70 \%$ and less than $65 \%$ of total wealth for non- and stockholders, respectively, with at least five percentage points of difference between the two categories. Risky

[^6]Table 3: ELSA Wave 8 summary statistics - financial respondent and non-respondent partners.

|  | All | Non-stockholders | Stockholders |
| :--- | :---: | :---: | :---: |
| financial respondent: male | $60,9 \%$ | $55,0 \%$ | $62,9 \%$ |
| age: respondent | 67,5 | 67,8 | 67,4 |
| age: non respondent | 67,1 | 67,1 | 67,1 |
| low edu: respondent | $28,9 \%$ | $46,6 \%$ | $22,9 \%$ |
| mid edu: respondent | $45,0 \%$ | $41,1 \%$ | $46,4 \%$ |
| high edu: respondent | $26,0 \%$ | $12,3 \%$ | $30,7 \%$ |
| low edu: non respondent | $33,1 \%$ | $47,2 \%$ | $28,3 \%$ |
| mid edu: non respondent | $45,4 \%$ | $43,6 \%$ | $46,1 \%$ |
| high edu: non respondent | $21,4 \%$ | $9,2 \%$ | $25,6 \%$ |
| general risk: respondent | 4,8 | 4,7 | 4,8 |
| general risk: non respondent | 4,5 | 4,4 | 4,6 |
| financial risk: respondent | 3,4 | 3,0 | 3,5 |
| financial risk: non respondent | 3,1 | 3,0 | 3,1 |

assets account for $10-15 \%$ of the household wealth, and their share is independent of income and rather stable among stockholders portfolio. This pattern is in line with the entry costs hypothesis presented in Section 2.3: comparing two households with the same optimal $\alpha$ and different financial wealth, the household with lower financial wealth invests a lower amount of money in risky assets, therefore the potential financial returns may not cover the fixed entry costs and the decison-makers decide to not hold risky assets.
These statistics are in line with the findings of Banks and Smith's (2000). They study the evolution of English households portfolio composition between 1980 and 2000, working with the Family Expenditure Survey (FES) and the Financial Research Survey (FRS) data. They show that housing and pensions funds account for the largest share of the household portfolio, with a progressive shift from housing towards financial assets over time. This shift was the consequence of tax-favoured products (TESSAs, replaced by ISAs in 1999) created by the government to try to encourage pension savings during the 90 s.

Table 4: Composition of total gross household wealth by stock market participation and household labour and pension income quartile.


[^7]
### 3.2 Weights

The model in Section 2 defines $\rho_{h}$ as a weighted household risk tolerance, where weights are the normalized bargaining powers of partners, $\gamma$ and $(1-\gamma)$ respectively. Measuring the bargaining powers of household members is one of the main issues of collective models. These powers may change relative to the kind of decision that the household is taking (e.g., the decision-making processes of financial choices may differ from other consumption choices like employment or childcare).
The first study that investigates the within-household allocation sharing rule and its determinants is Browning et al. (1994). They show that the allocation of resources is proportional to the relative share of household income of partners. In other words, the income pooling hypothesis, i.e. that it is the household total income that matter for the decision outcomes, is not consistent with the data. Their conclusion shows that it is the share of household income of each partner that affects the intrahousehold allocation of resources. Recently, Attanasio and Lechene (2014) study the within-household shift in bargaining power using data about the cash transfer program Progresa in rural Mexico. This state program generates a large variation in the wife's relative household income (about $20 \%$ of household total expenditures), explicitly changing the control of resources within the treated households. They show that this shift changes the balance of power within the couple, concluding that one of the determinants of the within-household bargaining power is the share of the current income of each partner.
I follow these results to construct a measure of partners bargaining powers. The couples of the sample manage their finances jointly and have a unique, shared household wealth. Within the household, one partner likely earns more than the other. Therefore, the income insurance that the high-income partner receives is lower than the one that he/she provides. To compensate this risk gap, the high-income partner has a higher control over household finances. Then, he/she can adjust the household risk, and the household portfolio, to cover the partial income insurance that the partner can not provide. Concluding, I assume that the bargaining power of each partner is proportional to their own share of household income, using the partners share of household income (labour, pension and state benefit income) as a proxy of the bargaining powers. The idea is that the higher is the wife/husband share of income, the higher is her/his control on the portfolio allocation decision.
Table 5 shows the income percentiles of ELSA partners by financial respondent and by gender $\sqrt{12}$, Male and financial respondent partners earn more. However, males median income is twice the median income of females, while this difference reduces among financial and non-financial respondents, even if it remains large and significant (the gap is about $35 \%$ of non-respondent median income).
Figures 3.2, 3.3 and 3.4 compare the box plots, the histograms and the kernel densities of wives and husbands weekly income distributions, measured in pounds. Both distributions are clearly non-normal and show positive skewness, with long right tails, and a concentration of the mass of the distribution on the left. Males income is higher (the husbands median income is two times the wives median income) and has a higher variability ( $s d_{m}=341.4 £$ against $s d_{f}=202.4 £$ of wives) and longer right tails. Wage differences may depend on individual circumstances (e.g.: number of dependent children, company size and type of occupation) and reduce since the introduction of the Equal Pay Act in 1975 in the UK, but still affect the English society.

[^8]Table 5: Partners weekly income percentiles. Income is the sum of employment and self-employment income, private and state pension income, and state benefit transfers.

|  |  |  |  |  |  |  |  | Income percentiles |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | individual <br> income $=0$ | $10 \%$ | $25 \%$ | $50 \%$ | $75 \%$ | $90 \%$ |  |  |  |  |  |  |  |
| male | 15 | $178 £$ | $258 £$ | $368 £$ | $543 £$ | $730 £$ |  |  |  |  |  |  |  |
| female | 73 | $51 £$ | $96 £$ | $173 £$ | $277 £$ | $407 £$ |  |  |  |  |  |  |  |
| financial respondent | 28 | $89 £$ | $179 £$ | $310 £$ | $487 £$ | $692 £$ |  |  |  |  |  |  |  |
| financial non-respondent | 60 | $69 £$ | $131 £$ | $231 £$ | $353 £$ | $533 £$ |  |  |  |  |  |  |  |

Table 6: Partners share of household income - percentiles. Income is the sum of employment and self-employment income, private and state pension income, and state benefit transfers.

|  | Share of hh income percentiles |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | share of hh <br> income $=0$ | $10 \%$ | $25 \%$ | $50 \%$ | $75 \%$ | $90 \%$ | share of hh <br> income $=1$ |
| male | 15 | $42.0 \%$ | $55.5 \%$ | $68.3 \%$ | $80.3 \%$ | $90.2 \%$ | 74 |
| financial respondent | 28 | $21.0 \%$ | $38.0 \%$ | $58.4 \%$ | $75.6 \%$ | $87.9 \%$ | 61 |

Figure 3.2: Female and Male income - boxplot.


Figure 3.3: Female and Male income - histogram.


Figure 3.4: Female and Male partners income - kernel densities.


Table 6 shows the percentile of share of household income $(\gamma)$ of financial respondent and males partners. On average, the male partner earns the $66.5 \%$ (sd: 19.9\%) of household labour and pension income, while the financial responding partner share is $56.3 \%$ (sd: $25.2 \%$ ). The fact that the financial respondent has $10 \%$ less power depends mainly on the share of females financial respondents ( $40 \%$ circa), which earn a lower income than their partners, as Figure 3.2, 3.3 and 3.4 and Table 5 show.

## 4 Results

This Section analyses empirically the implication of the theoretical model described in Section 2 , First, I compare two econometric models that estimate household portfolio allocation using the general and the financial risk tolerance, respectively. This first step drives the choice of the baseline model, which uses the financial measure. Then, I discuss the results of the estimates, focusing on the determinants of household stock market participation and household share of wealth allocated in risky assets. Finally, I provide evidence that the collective model fits significantly better the data compared to the standard unitary model generally used in the literature.
I present a reduced-form analysis, that studies the effect of the collective risk tolerance on household portfolio allocation conditioning on a set of household demographics. The estimation of a structural model would allow studying how the demographics determine the household risk preferences and the role of the unobserved heterogeneity component of risk tolerance (see Section 2.2), however, this goes beyond the scope of this paper and is left for future work.
The empirical analysis does not consider distribution factors that are relevant elements of the collective models. Distribution factors are defined as exogenous conditions that may affect the bargaining power of the spouses without altering their preferences. Some examples are divorce law or other policies that explicitly change the distribution of resources within the household (e.g., see Attanasio and Lechene (2014)). Due to data limitations, it is not possible to identify proper distribution factors and study their effects in this analysis.

### 4.1 Truncation correction method

The estimates of the household optimal portfolio allocation present a problem of incidental truncation: the explanatory variables are always observed, while the dependent variable (the share of household wealth invested in risky assets, $\alpha$ ) is available only for a subset of the population. In other words, I observe $\alpha$ only for those households that decide to invest in risky assets. Therefore, the rule determining whether $\alpha$ is observed or not does not depend directly on the outcome of $\alpha$ itself. Concluding, the truncation of the dependent variable is incidental because it depends on household decision to participate in the stock market.

I use the approach proposed by Heckman (1979) to estimate the effects of risk preferences on household portfolio allocation. This method allows the correction of bias from non-randomly selected samples or incidentally truncated dependent variables. First, it estimates the probability of observing the dependent variable using a probit model (selection equation) and then includes these results in the linear OLS estimation of the dependent variable (outcome equation). In other words, the second stage corrects for non-random selection by incorporating a transformation of the predicted probabilities of observing the dependent variable as an additional explanatory variable. This is the so-called Heckman two-step procedure.
In this analysis, the two dependent variables of the Heckman procedure are household stock market participation (selection equation) and the share of wealth invested in risky assets (outcome equation). In other words, the first stage estimates the probability $P(\alpha>0)$, while the second stage estimates $\alpha$.

### 4.2 Exclusion restriction

An exclusion restriction is required for non-parametric identification. There must be at least one variable that appears with a non-zero coefficient in the selection equation but does not appear in the outcome equation: this variable is essentially an instrument.
In this case, the selection equation includes age and age squared of the male partner, a dummy for large age difference in the couple (more than 10 years), household income and net total wealth quartiles, job market participation, a sickness index ${ }^{[3]}$, education, risk tolerance, patience scores and numeracy score of both partners. The outcome equation does not include partners' numeracy scores, which serves as exclusion restriction.
Numeracy approximates the individual cognitive skills. I assume that agents with low numeracy need more time to improve their financial knowledge, increasing the stock market fixed entry costs. These higher costs decrease the potential returns of their investments. Then, those households with low numeracy may decide to not hold stocks. This assumption implies that all the households below a minimum threshold of numeracy do not participate in the stock market, while all the stockholders are above that threshold. Thus, numeracy affects household stock market participation, but the heterogeneity in $\alpha$ of those who participate does not depend on cognitive skills. Numeracy is based on a set of questions about simple math exercises, like computing percentages, fractions, additions and subtractions. I use the 5 questions that ask to compute a sequence of subtractions: respondents have to subtract 7 from 100, and then 7 from the previous result and so on, five times. I compute individual numeracy score as the sum of correct answers of the respondent. I aggregate the scores

[^9]from 2 to 4 , obtaining a total of 4 possible categories for each partner: no numeracy skills ( 0 correct answers), low numeracy ( 1 correct answer), medium numeracy ( 2 to 4 correct answers) and high numeracy ( 5 correct answers) Tables 7 shows the construction and the distribution of financial literacy. On average, the numeracy score is higher among stockholders and males and financial respondents, i.e., those with higher bargaining power, show higher numeracy skills than females and non-financial respondents, respectively.

Table 7: Financial literacy: partners by gender

| Male fin. literacy |  |  |  |
| :---: | :---: | :---: | :---: |
| numeracy | Freq. | Percent | Cum. |
| 1 | 26 | $1,8 \%$ | $1,8 \%$ |
| 2 | 132 | $9,4 \%$ | $11,2 \%$ |
| 3 | 258 | $18,3 \%$ | $29,6 \%$ |
| 4 | 523 | $37,2 \%$ | $66,8 \%$ |
| 5 | 467 | $33,2 \%$ | $100,0 \%$ |
| Total | 1,406 | 100.00 | - |


| Female fin. literacy |  |  |  |
| :---: | :---: | :---: | :---: |
| numeracy Freq. Percent Cum. <br> 1 114 $8,1 \%$ $8,1 \%$ <br> 2 279 $19,8 \%$ $28,0 \%$ <br> 3 352 $25,0 \%$ $53,0 \%$ <br> 4 451 $32,1 \%$ $85,1 \%$ <br> 5 210 $14,9 \%$ $100,0 \%$ <br> Total 1,406 100.00 - |  |  |  |

### 4.3 Empirical estimates

This Section presents the empirical results of the paper. First, I focus on the selection of the collective model that serves as a benchmark through the analysis, then, I study the determinants of the household portfolio allocation and, last, I compare the collective and the unitary approach. The main results presented in this paper distinguish partner characteristics (demographics and risk) by gender. In Appendix Section CI will provide additional robustness checks using financial respondent and non-respondent partner demographics.

### 4.3.1 Baseline model selection

The ELSA survey provides two self-assessed measures of individual risk preferences. The former asks about risk preferences in general, while the latter asks explicitely about risk preferences in financial contexts (spending, savings).
Table 8 compares the outcome equation of the Heckman estimates of two specification that differ because of the type of risk tolerance used: Column (1) uses the financial risk tolerance, and Column (2) uses the general risk tolerance. Note that both risk measures are at the household level. Then, they are the weighted sum of partners risk preferences, where the weights are the share of household income of each spouse.
I use maximum likelihood estimation, which provides two evaluation metrics of the goodness of fit of the models, the information criteria Akaike (AIC) and Schwarz (BIC). Results are qualitatively the same and, based on AIC and BIC of Table 8, Column (1), i.e. the model that uses financial risk preferences, has to be preferred. Therefore, in what follows I focus on financial risk preferences and use Column (1) Table 8 as the baseline model.

### 4.3.2 Household portfolio allocation

This Section discusses the results of the empirical estimates that study the determinants of the household portfolio allocation decision.

Table 8: Heckman outcome equation: household share of net financial wealth allocated in risky assets. General vs financial risk tolerance measure. Partners characteristics by gender.

| share of risky assets | Financial <br> risk measure <br> $(1)$ | General <br> risk measure <br> $(2)$ |
| :--- | :---: | :---: |
| demographics $\mathbb{m}^{\bar{a}}$ | $*$ | $*$ |
| financial risk: hh | $0.0354^{* * *}$ |  |
|  | $(0.0053)$ |  |
| general risk: hh |  | $0.0120^{* *}$ |
|  |  | $(0.0051)$ |
| Constant | -0.3596 | -0.1872 |
|  | $(0.6821)$ | $(0.6866)$ |
| AIC | $\mathbf{1 9 9 3 . 2}$ | $\mathbf{2 0 2 8 . 4}$ |
| BIC | $\mathbf{2 2 2 4}$ | $\mathbf{2 2 5 9}$ |
| Number of observation | 1,406 | 1,406 |

${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners (more than 10 years), dummies of income and wealth quartiles, job market participation of partners, education of partners and the sickness index of partners.

I estimate two different collective specifications that capture the effects of risk preferences of both partners on household portfolio allocation. Table 9 presents the results of the Heckman first stage: Column (1) uses the measure of husband and wife (financial) risk tolerance separately, while Column (2) uses the household weighted risk measures.
Income and wealth strongly affect the probability of stock market participation: income effects are stable across its quartiles, while wealth effects are increasing in magnitude. These findings are in line with the entry costs assumption in Section 2.3; the higher is the household wealth, the lower is the impact of lump-sum costs on household finances and the probability of being/becoming a stockholder increases. The numeracy of both partners is significant and have a positive effect on the stock market participation ${ }^{14}$. Male numeracy effects are increasing, while female numeracy does not show a similar pattern. This difference might be a consequence of the higher bargaining power of males: when the husband has high cognitive skills, he increases the probability of household stock market participation because of the higher (average) influence. Male job market participation decreases the household probability of being a stockholder. Husbands may hold the "last say" on portfolio decision because of their higher (average) bargaining power. When the husband is working, he has less time to collect information about financial markets and then may prefer to not invest in risky assets. Moreover, workers contributing to a Defined Contribution pension plan receive a lump-sum payment which amounts to about $25 \%$ of their pension pot when transiting to retirement. Therefore, retired individuals have higher liquidity to invest in the stock market at the beginning of retirement. Male education has positive and significant effects, as largely documented in the literature. Last, there are no significant effects of financial risk tolerance on stock market participation, in line with the theoretical model. Indeed, risk tolerance must affect the share of wealth allocated in risky assets, but not participation, which depends on wealth, stock market entry costs, and numeracy.
The outcome equation, in Table 10, estimates a linear model where the dependent variable is $\alpha$ (share of household net financial wealth allocated in risky assets) and the regressors are the same of

[^10]the probit estimates of the first stage. However, the outcome equation excludes partners numeracy scores and includes the inverse Mills ratio of the first stage. The structure of Table 10 is the same of Table 9.
$\alpha$ increases in net financial wealth, with wealthier households that invest a larger share of their finances in risky assets. The female partner high education increases the portfolio share allocated in risky investments, while the male education has no effects. Notice that the opposite is true in the first stage. This effect might be due to partners sorting into marriage based on education. Therefore, couples participating in the stock market tend to have higher education on average. Once the stockholders are considered, what seems to matter is the high education level of the wife, because a woman with a college degree is rarer than a man with the same title.
Risk tolerance positively affects the share of wealth allocated in risky assets in both cases, when the partners' or the household risk preferences is considered. The difference is in the marginal effects: male and female partner risk show similar coefficients, while the weighted household risk tolerance impact is twice the risk tolerance of the two partners.
The results are in line with the collective portfolio decision model, which describes the household financial decision process as a two-step procedure. In the first step, partners choose the shared degree of risk tolerance and then decide about stock market participation. In the second step, the choice concerns the optimal share of wealth allocated in risky assets, if any. Therefore, the household takes decisions as a system of individuals that combine their preferences, and not as a single unit. The specifications in Table 9 and Table 10 follow this idea: each of them includes the preferences of all the household decision-makers, combined in a unique (weighted) measure or not, and these preferences affect the household portfolio allocation. Last, the inverse Mills ratio presented in Table 10 is non-significant. However, as I will discuss in the robustness section (Section Cin the Appendix), controlling for selection remains important in this context.

### 4.3.3 Collective vs unitary approach

This Section aims to assess whether the collective approach proposed in this paper fits the data significantly better than the standard unitary approach generally used in the literature to study the household portfolio allocation.
As stated in the Introduction, the unitary model describes the household as a single decision-making unit that solves the utility maximization problem with well-defined preferences. Empirically, it is common practice to proxy the household behaviour using the husband or the head of the household. In this paper, I represent the unitary household with the risk preferences of the husband. In the Appendix, I present the result obtained using the household head as a proxy of the unitary household, i.e. with the preferences of the partner who hold the last say on the decision (e.g.: Bertocchi et al. (2014)). In this context, I proxy the household head using the financial respondent of the interview, i.e. the partner that answers to the Income \& Asset section of the survey.

I compare three specifications, two collective models and one unitary model, using the information criteria and the likelihood ratio test. The selection and the outcome equation of each specification are jointly estimated using maximum likelihood, which provides the values of the AIC and BIC to compare the goodness of fit of the models. Last, the likelihood ratio test is constructed by considering the unitary model as a special case of the collective models. Then, the unitary model is the nested (or reduced) model of the test.

Table 9: Household portfolio determinants: household probability of holding risky assets. Selection equation. Partners characteristics by gender.

| participation | Financial risk |  |
| :---: | :---: | :---: |
|  | Individual (1) | Baseline (2) |
| fin. literacy: male | $0.1900^{* * *}$ | $0.1886^{* * *}$ |
|  | (0.0403) | (0.0402) |
| fin. literacy: female | $0.0942^{* * *}$ | $\begin{aligned} & 0.0961^{* * *} \\ & (00358) \end{aligned}$ |
| age: male | $0.1407^{* *}$ | $0.1400{ }^{* *}$ |
|  | (0.0646) | (0.0644) |
| age ${ }^{2}$ : male | -0.0009** | -0.0009** |
|  | (0.0005) | (0.0005) |
| age diff $>10$ | 0.0720 | 0.0529 |
|  | (0.1649) | (0.1646) |
| $2^{\text {nd }}$ income quartile | $\begin{gathered} 0.2743^{* * *} \\ (0.1039) \end{gathered}$ | $\begin{gathered} 0.2751^{* * *} \\ (0.1037) \end{gathered}$ |
| $3^{\text {rd }}$ income quartile | $0.4823^{* * *}$ | $0.4736^{* * *}$ |
|  | (0.1138) | (0.1136) |
| $4^{\text {th }}$ income quartile | 0.5662*** | $0.5722^{* * *}$ |
|  | (0.1276) | (0.1275) |
| in work: male | -0.2289** | $-0.2317 * *$ |
|  | (0.1079) | (0.1078) |
| in work: female | $-0.1554$ | $-0.1634$ |
| mid education: male | $0.2513^{* * *}$ | $0.2518^{* * *}$ |
|  | (0.0956) | (0.0955) |
| high education: male | $0.4337^{* * *}$ | $0.4229^{* * *}$ |
|  | (0.1324) | (0.1321) |
| mid education: female | 0.0860 | 0.0849 |
|  | (0.0936) | (0.0935) |
| high education: female | 0.2628* | $0.2615^{*}$ |
|  | (0.1419) | ${ }^{(0.1416)}$ |
| sickness: male | -0.0616*** | -0.0607*** |
|  | (0.0220) | (0.0219) |
| sickness: female | -0.0339 | -0.0368* |
|  | (0.0207) | (0.0207) |
| financial risk: male | $\begin{aligned} & 0.0304^{*} \\ & (0.0167) \end{aligned}$ |  |
| financial risk: female | -0.0129 |  |
|  | (0.0171) |  |
| financial patience: male | -0.0177 |  |
|  | (0.0190) |  |
| financial patience: female | 0.0189 |  |
| financial risk: hh |  |  |
|  |  | (0.0199) |
| financial patience: hh |  | -0.0131 |
| Constant | -6.1515*** | -6.0119*** |
|  | (2.2622) | (2.2584) |
| Number of observations | 1,406 | 1,406 |

In line with the selected baseline specification of Section 4.3.1, the three estimated models use the financial risk preferences of households and partners. The three specifications share the same first-stage selection equation, which uses the numeracy of both partners as exclusion restriction and includes the household demographics and the risk tolerances of each partner, separately. On the other hand, the specifications have three different outcome equations, whose estimation results are shown in Table 11. In particular, the first collective model (Column (1)) uses the risk tolerance of each partner separately (wife and husband), the second collective model (Column (2)) includes the household weighted risk tolerance and, additionally, the risk preference of the husband, while the unitary model (Column (3)) uses only the husband risk preferences (as a proxy of the household preferences). Therefore, the risk tolerance of the husband appears in each outcome equation, such

Table 10: Household portfolio determinants: household share of net financial wealth allocated in risky assets. Outcome equation. Partners characteristics by gender.

| share of risky assets | Financial risk |  |
| :---: | :---: | :---: |
|  | Individual (1) | $\begin{aligned} & \text { Baseline } \\ & \text { (2) } \end{aligned}$ |
| age: male | 0.0088 | 0.0078 |
|  | (0.0204) | (0.0205) |
| age ${ }^{2}$ : male | -0.0001 | -0.0000 |
|  | (0.0001) | (0.0001) |
| age diff $>10$ | $\begin{gathered} -0.0051 \\ (0.0436) \end{gathered}$ | $\begin{gathered} -0.0141 \\ (0.0438) \end{gathered}$ |
| $2^{\text {nd }}$ income quartile | -0.0354 | -0.0395 |
| $3^{\text {rd }}$ income quartile | (0.0327) | (0.0328) |
|  | -0.0331 | -0.0370 |
|  | (0.0370) | (0.0371) |
| $4^{\text {th }}$ income quartile | $-0.0345$ | -0.0417 |
|  | (0.0404) | (0.0408) |
| in work: male | -0.0165 | -0.0219 |
|  | (0.0292) | (0.0294) |
| in work: female | $\begin{gathered} -0.0175 \\ 0 \end{gathered}$ | $\begin{gathered} -0.0051 \\ \hline \end{gathered}$ |
| mid education: male | -0.0057 | -0.0066 |
|  | (0.0310) | (0.0312) |
| high education: male | $-0.0050$ | $-0.0081$ |
| mid education: female | 0.0197 | 0.0160 |
|  | (0.0271) | (0.0273) |
| high education: female | $0.0997^{* * *}$ | $0.1019^{* * *}$ |
|  | (0.0365) | (0.0368) |
| sickness: male | 0.0023 | 0.0033 |
| sickness: female | ${ }_{-0}(0.0037)$ | (0.0037) |
|  | $\begin{aligned} & -0.0038 \\ & (0.0064) \end{aligned}$ | $\begin{gathered} -0.0032 \\ (0.0064) \end{gathered}$ |
| financial risk: male | $0.0176^{* * *}$ |  |
|  | (0.0047) |  |
| financial risk: female | $\begin{gathered} 0.0208^{* * *} \\ (00047) \end{gathered}$ |  |
| financial patience: male | 0.0015 |  |
|  | (0.0050) |  |
| financial patience: female | 0.0008 |  |
| financial risk: hh |  | $0.0340^{* * *}$ |
|  |  | (0.0055) |
| financial patience: hh |  | 0.0026 |
|  |  | (0.0062) |
| Constant | $-0.0145$ | 0.0451 |
|  | (0.7569) | (0.7587) |
| ${ }_{\lambda}^{\text {Inverse Mills ratio }}$ |  |  |
|  | -0.1943** | -0.2046** |
|  | (0.0941) | (0.0951) |
| Number of observations | 1,406 | 1,406 |
| Selection | 1,048 | 1,048 |

that the unitary model (Column (3)) becomes a nested model of the two collective specifications (Columns (1) and (2)).
I compare the three specifications using the likelihood ratio test, testing if the risk preferences of wives add information to the collective models and matter in the household portfolio allocation process.

Table 11 shows the likelihood ratio test results between Column (1) and Column (3) and between Column (2) and Column (3), i.e. between the two collective models and the unitary model. The test rejects the null hypothesis in both cases. Thus, the additional variable used in Column (1) and (2), i.e. the risk tolerance of the wives, has an important role in explaining household portfolio allocation.

Last, Table 12 reports the baseline model selected in Section 4.3 .1 and compares the AIC and BIC information criteria of the collective and the unitary model. I conclude that the collective approach proposed in this paper fits significantly better the data than the standard unitary approach, where the information about the second partner risk preferences are missing.

Table 11: Collective vs unitary approach - likelihood ratio test. Outcome equation: household share of net financial wealth allocated in risky assets. Partners characteristics by gender.

| share of risky assets | Collective models |  | Unitary <br> (3) |
| :---: | :---: | :---: | :---: |
|  | Partner risk <br> (1) | Household risk (2) |  |
| demographic $\square^{\text {a }}$ | * | * | * |
| financial risk: hh |  | $\begin{gathered} 0.0463^{* * *} \\ (0.0102) \end{gathered}$ |  |
| financial risk: male | $\begin{gathered} 0.0188^{* * *} \\ (0.0044) \end{gathered}$ | $\begin{aligned} & -0.0105 \\ & (0.0085) \end{aligned}$ | $\begin{gathered} 0.0222^{* * *} \\ (0.0044) \end{gathered}$ |
| financial risk: female | $\begin{gathered} 0.0206^{* * *} \\ (0.0046) \end{gathered}$ |  |  |
| Constant | $\begin{array}{r} -0.3431 \\ (0.6794) \\ \hline \end{array}$ | $\begin{array}{r} -0.3177 \\ (0.6789) \\ \hline \end{array}$ | $\begin{gathered} -0.2274 \\ (0.6847) \end{gathered}$ |
| likelihood ratio test p -value | $\begin{gathered} \mathrm{Col}(1) \text { and (3) } \\ 0.000 \end{gathered}$ | $\begin{gathered} \mathrm{Col}(1) \text { and (3) } \\ 0.000 \end{gathered}$ |  |
| Number of observations | 1,406 | 1,406 | 1,406 |

[^11]Table 12: Collective vs unitary approach - information criteria. Heckman outcome equation: household share of net financial wealth allocated in risky assets. Partners characteristics by gender.

| share of risky assets | Collective model <br> $(1)$ | Unitary model <br> $(2)$ |
| :--- | :---: | :---: |
| demographics ${ }^{\bar{a}}$ | $*$ | $*$ |
| financial risk: hh | $0.0352^{* * *}$ |  |
|  | $(0.0053)$ |  |
| financial risk: male |  | $0.0220^{* * *}$ |
|  |  | $(0.0044)$ |
| Constant | -0.2980 | -0.2202 |
|  | $(0.6785)$ | $(0.6825)$ |
| AIC | $\mathbf{1 9 8 7 . 9}$ | $\mathbf{2 0 0 9 . 9}$ |
| BIC | $\mathbf{2 1 8 7}$ | $\mathbf{2 2 0 4}$ |
| Number of observations | 1,406 | 1,406 |

[^12]
### 4.3.4 Household portfolio allocation - under 65 years

This Section presents the results of previous Sections using households where at least one partners is aged less than 65 years. Older agents have a high probability of being retired and may change their investment preferences, financial behaviours and decision-making process. This subsample is composed by 676 households and the stock market participation rate is $73.4 \%$, in line with the overall sample.
Table 13 and 14 compares the collective and the unitary model using the likelihood ratio test and the Bayesian information criteria, following the analysis explained in Section 4.3.3.

Table 13: Under 65 years: collective vs unitary approach - likelihood ratio test. Outcome equation: household share of net financial wealth allocated in risky assets. Partners characteristics by gender.

| Under 65 years |  |  |  |
| :---: | :---: | :---: | :---: |
|  | Collective models |  | Unitary <br> (3) |
| share of risky assets | Partner risk <br> (1) | Household risk <br> (2) |  |
| demographic $\square^{\square}$ | * | * | * |
| financial risk: hh |  | $\begin{gathered} 0.0409^{* * *} \\ (0.0131) \end{gathered}$ |  |
| financial risk: male | $\begin{gathered} 0.0220^{* * *} \\ (0.0068) \end{gathered}$ | $\begin{gathered} -0.0024 \\ (0.0115) \end{gathered}$ | $\begin{gathered} 0.0269^{* * *} \\ (0.0067) \end{gathered}$ |
| financial risk: female | $\begin{gathered} 0.0232^{* * *} \\ (0.0066) \end{gathered}$ |  |  |
| Constant | $\begin{aligned} & 3.5133^{* *} \\ & (1.4808) \\ & \hline \end{aligned}$ | $\begin{aligned} & 3.5757^{* *} \\ & (1.4885) \\ & \hline \end{aligned}$ | $\begin{aligned} & 3.4335^{* *} \\ & (1.4973) \\ & \hline \end{aligned}$ |
| likelihood ratio test | Col (1) and (3) | Col (2) and (3) |  |
| p-value | 0.0019 | 0.0079 |  |
| Number of observations | 676 | 676 | 676 |

[^13]Table 14: Under 65 years: collective vs unitary approach - information criteria. Outcome equation: household share of net financial wealth allocated in risky assets. Partners characteristics by gender.

| Under 65 years |  |  |
| :--- | :---: | :---: |
| share of risky assets | Collective model <br> $(1)$ | Unitary model <br> $(2)$ |
| demographics $]^{a}$ | $*$ | $*$ |
| financial risk: hh | $0.0380^{* * *}$ |  |
| financial risk: male | $(0.0077)$ |  |
| Constant |  | $0.0266^{* * *}$ |
|  | $3.7726^{* *}$ | $(0.0067)$ |
| AIC | $(1.4885)$ | $(1.4958)$ |
| BIC | $\mathbf{1 0 1 1}$ | $\mathbf{1 0 1 8}$ |
| Number of observations | $\mathbf{1 1 8 3}$ | $\mathbf{1 1 9 0}$ |

${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the health index of partners.

The table show that the results are in line with previous findings, even if somewhat attenuated. In Section C. 4 of the Appendix I perform the analysis distinguishing partners characteristics by financial and non-financial respondent, as a robustness check. These estimates confirm the results of Section 4, using the financial respondent partner risk preferences to approximate the unitary household.

### 4.4 Bargaining power estimates

The tables presented in Section 4.3 show the estimate of two models that differs in how the risk preferences of the household decision-makers are included. Formally:

$$
\begin{equation*}
Y=\alpha_{0}+\alpha_{1} \cdot \rho_{m}+\alpha_{2} \cdot \rho_{f}+\alpha^{\prime} X+\epsilon \tag{16}
\end{equation*}
$$

Figure 4.5: Estimated and approximated average bargaining power by household subgroup. Dark grey bars represent the $\gamma$ computed as described in Equation 20, and the light grey bars represent the average $\gamma$ of the subgroups.


$$
\begin{equation*}
Y=\beta_{0}+\beta_{1} \cdot \rho_{h h}+\beta^{\prime} X+\epsilon \tag{17}
\end{equation*}
$$

where $\rho_{m}, \rho_{f}$ and $\rho_{h h}$ represent the male, female and household (weighted) risk tolerance, respectively. Considering that $r_{h h}$ is: $\rho_{h h}=\gamma \cdot \rho_{m}+(1-\gamma) \cdot r_{f}$, I can rewrite Equation 17 as:

$$
\begin{align*}
& Y=\beta_{0}+\beta_{1} \cdot\left(\gamma \cdot r_{m}+(1-\gamma) \cdot r_{f}\right)+\beta^{\prime} X+\epsilon \\
& Y=\beta_{0}+\beta_{1} \gamma \cdot r_{m}+\beta_{1}(1-\gamma) \cdot r_{f}+\beta^{\prime} X+\epsilon \tag{18}
\end{align*}
$$

Thus, from the equivalence between Equation 16 and Equation 18 I obtain the following:

$$
\begin{cases}\alpha_{1} & =\beta_{1} \gamma  \tag{19}\\ \alpha_{2} & =\beta_{1}(1-\gamma)\end{cases}
$$

rearranging, I obtain:

$$
\begin{equation*}
\gamma=\frac{\alpha}{1+\alpha} \tag{20}
\end{equation*}
$$

where $\alpha=\frac{\alpha_{1}}{\alpha_{2}}$. In other words, it is possible to estimate the average bargaining power ( $\gamma$ ) of households directly using the results of the empirical estimates.

The result of this process is, as stated above, the average value of $\gamma$ defined in Section 3.2. Figure 4.5 shows the various $\gamma$ computed for different household subgroups of the sample: gender of financial respondent, presence of at least one household member aged below 65 years, and presence of at least one employed/self-employed member.
The figure compares the outcomes of the strategy proposed in Section 3.2 to approximate the household bargaining powers and the $\gamma$ computed from the empirical estimates. The approximation proposed in Section 3.2 tends to overestimate the average $\gamma$ estimated following Equation 20 , However, the $95 \%$ confidence intervals constructed using bootstrap always include the measured bargaining power. Thus, the methodology used in this paper to approximate the bargaining power is consistent.

## 5 Conclusion

This paper studies the role of partners risk preferences in the household portfolio choice process. I develop a theoretical model that describes the household portfolio allocation decision following the collective approach proposed by Chiappori (1988). This approach considers the household as a group of agents which combine their preferences through a bargaining process. Formally, the general model assumes that the decision-making process is a weighted combination of individual behaviours, where the weights are the bargaining power of each household's member. Assuming exponential utility, I show how the optimal portfolio allocation depends on a weighted average of partners risk preferences, fixed stock market entry costs, and stock market returns.

I use the ELSA survey to study the effects of risk preferences on household portfolio allocation, using the share of household income as a proxy of the bargaining powers. The empirical estimate relies on the Heckman selection model, which corrects for the bias produced by non-randomly selected samples derived by stock market participation. The exclusion restriction relies on partners' financial literacy. Low financial knowledge increases the stock market fixed entry costs because the household needs more time to learn about the stock market. Then, the partners may decide to not hold risky assets because of lower potential returns. The estimates show that stock market participation increases in household income, household wealth, partners numeracy and education, while risk tolerance has no effects. On the other hand, a higher risk tolerance increases the share of wealth allocated in risky assets, once the household is a stockholder.
Last, I compare the collective and the unitary approach estimating three different specifications with maximum likelihood. I use the likelihood ratio test and Akaike's and Schwarz's Bayesian information criteria to compare their goodness of fit. The results show that the collective approach performs significantly better than the unitary one. Thus, it is relevant to account for the preferences of both spouses when studying household portfolio allocation.

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## Additional Material Households risk preferences and portfolio allocation: a collective approach

Francesco Maura

## A Theoretical model

## A. 1 CARA and mean-variance utility

Household wealth $w$ can be invested in two types of assets: a risk-free asset with constant return $(r)$ and a risky asset with return $\tilde{r}=r+\tilde{s}$ where $\tilde{s} \sim N\left(\mu_{s}, \sigma_{s}^{2}\right)$. The household (or agent) $i$ maximizes its own utility $u_{i}$, choosing the optimal share of wealth allocated in risky asset, $\alpha$. Wealth after the investments is:

$$
W_{i}=\left(w_{h}-\alpha\right)(1+r)+w_{h} \alpha(1+\tilde{r})=w_{h}(1+r+\alpha \tilde{s})
$$

Assuming CARA utility with risk aversion parameter $\rho_{i}$, we have:

$$
u_{i}\left(W_{h}\right)=-e^{-\rho_{i} W_{h}}
$$

Because risky asset returns are normally distributed, the value function $V_{i}(\alpha)$ can be written as:

$$
\begin{align*}
\max _{\alpha} V_{i}(\alpha) & =\max _{\alpha} E\left[u_{i}\left(W_{h}\right)\right]=\max _{\alpha} E\left[-e^{-\rho_{h} W_{h}}\right] \\
& =\max _{\alpha}-E\left[e^{-\rho_{i}\left(w_{h}[(1+r)+\alpha \tilde{s}]\right.}\right]=\max _{\alpha}-\ln \left(E\left[e^{-\rho_{i} w_{h}[(1+r)+\alpha \tilde{s}]}\right]\right) \\
& =\max _{\alpha}-\left(E\left[-\rho_{i} w_{h}[(1+r)+\alpha \tilde{s}]\right]+\frac{1}{2} \operatorname{Var}\left[\rho_{h} w_{h}[(1+r)+\alpha \tilde{s}]\right]\right)  \tag{1}\\
& =\max _{\alpha} \rho_{i} w_{h}\left[(1+r)+\alpha \mu_{s}\right]-\frac{1}{2}\left[\alpha^{2} \sigma_{s}^{2} \rho_{i}^{2} w_{h}^{2}\right] \\
& =\max _{\alpha} \rho_{i} w_{h}\left[(1+r)+\alpha \mu_{s}-\frac{1}{2} \alpha^{2} \sigma_{s}^{2} \rho_{i} w_{h}\right]
\end{align*}
$$

## A. 2 Household utility as a weighted sum of decision makers utilities

I propose a second model of household portfolio allocation. With respect to Section 2.1, the household maximizes a weighted sum of the two decision-makers utilities. I assume a CARA utility function $u_{i}\left(W_{h}\right)$ for each decision-maker ( $a$ and $b$ ), and egoistic preferences of agents (i.e.: their utility does not depends on partners utility). Wealth can be invested in two types of assets: a risk-free asset with constant return $(r)$ and a risky asset with return $(1+\tilde{r})=(1+r+\tilde{s})$ where $\tilde{s} \sim N\left(\mu_{s}, \sigma_{s}^{2}\right)$. Therefore, the problem of agent $i$ is to maximize his/her own utility $u_{i}$, choosing the optimal share of wealth allocated in risky asset, $\alpha$.
Wealth is defined as

$$
W_{h}=w_{h}(1+r+\alpha \tilde{s})
$$

and $u_{i}$ becomes:

$$
u_{i}\left(W_{h}\right)=-e^{-\rho_{i} W_{h}}
$$

Individual $i$ maximizes:

$$
\begin{align*}
\max _{\alpha} V_{i}(\alpha) & =\max _{\alpha} E\left[u_{i}\left(W_{h}\right)\right]=\max _{\alpha} E\left[-e^{-\rho_{h} W_{h}}\right] \\
& =\max _{\alpha} \rho_{i} w_{h}\left[(1+r)+\alpha \mu_{s}-\frac{1}{2} \alpha^{2} \sigma_{s}^{2} \rho_{i} w_{h}\right] \tag{2}
\end{align*}
$$

while household utility $u_{h}\left(W_{h}\right)$ is:

$$
\begin{equation*}
u_{h}\left(W_{h}\right)=\mu_{a} u_{a}\left(W_{h}\right)+\mu_{b} u_{b}\left(W_{h}\right) \tag{3}
\end{equation*}
$$

where the weights $\left(\mu_{a}, \mu_{b}\right)$ are the bargaining powers of the two individuals. I normalize the weights as follow:

$$
\begin{equation*}
\gamma=\frac{\mu_{a}}{\mu_{a}+\mu_{b}} \quad(1-\gamma)=\frac{\mu_{b}}{\mu_{a}+\mu_{b}} \tag{4}
\end{equation*}
$$

such that $\gamma \in[0,1]$ represents agent's $a$ bargaining power. Thus, the household maximization problem in Equation 3 becomes:

$$
\begin{equation*}
u_{h}\left(W_{h}\right)=\gamma u_{a}\left(W_{h}\right)+(1-\gamma) u_{b}\left(W_{h}\right) \tag{5}
\end{equation*}
$$

In terms of the value function $V_{h}(\alpha)$, the program becomes:

$$
\begin{align*}
\max _{\alpha} V_{h}(\alpha) & =\max _{\alpha} E\left[u_{h}\left(W_{h}\right)\right] \\
& =\max _{\alpha} E\left[\gamma u_{a}\left(W_{h}\right)+(1-\gamma) u_{b}\left(W_{h}\right)\right] \\
& =\max _{\alpha} \gamma E\left[u_{a}\left(W_{h}\right)\right]+(1-\gamma) E\left[u_{b}\left(W_{h}\right)\right]  \tag{6}\\
& =\max _{\alpha} \gamma \rho_{a} w_{h}\left[(1+r)+\alpha \mu_{s}-\frac{1}{2} \alpha^{2} \sigma_{s}^{2} w \rho_{a}\right]+ \\
& +(1-\gamma) \rho_{b} w_{h}\left[(1+r)+\alpha \mu_{s}-\frac{1}{2} \alpha^{2} \sigma_{s}^{2} w \rho_{b}\right]
\end{align*}
$$

Then, solving the first order conditions for $\alpha$ :

$$
\begin{equation*}
\alpha=\frac{\mu_{s}}{\sigma_{s}^{2} w} \cdot \frac{\rho_{h}}{\phi} \tag{7}
\end{equation*}
$$

where $\rho_{h}=\gamma \rho_{a}+(1-\gamma) \rho_{b}$ and $\phi=\gamma \rho_{a}^{2}+(1-\gamma) \rho_{b}^{2}$.

## B ELSA dataset

Table 1: Sample selection

| Selection | Decrease in observation |
| :--- | :---: |
| Initial sample (individuals) | 8445 |
| Individuals in a couple with joint finances | 4965 |
| Couple with joint finances (both partners) | 2325 |
| Couple with both partners younger than 90s | 2304 |
| Couple with both partners self-reported risk valid answers | 1582 |
| Couple with positive income | 1577 |
| Couple with non-negative net wealth | 1465 |
| Couple with risky share of wealth $\in[0,1]$ | 1445 |
| sickness index missing values | 1442 |

## C Results - robustness check

## C. 1 Under 65 years - portfolio determinants

Table 2: Under 65 years: baseline model selection: financial and general risk tolerance. Heckman selection equation: household probability of holding risky assets. Partners characteristics by gender.

| Under 65 years |  |  |
| :---: | :---: | :---: |
| participation | $\begin{aligned} & \text { Financial } \\ & \text { risk measure } \\ & \text { (1) } \end{aligned}$ | General risk measure (2) |
| low numeracy: male | $\begin{gathered} 0.3287 \\ (0.3897) \end{gathered}$ | $\begin{gathered} 0.3405 \\ (0.3881) \end{gathered}$ |
| mid numeracy: male | $\begin{gathered} 0.4603 \\ (0.4008) \end{gathered}$ | $\begin{gathered} 0.4791 \\ (0.3997) \end{gathered}$ |
| high numeracy: male | $\begin{gathered} 0.5796 \\ (0.3686) \end{gathered}$ | $\begin{gathered} 0.5837 \\ (0.3670) \end{gathered}$ |
| low numeracy: female | $\begin{gathered} 0.8346^{* * *} \\ (0.2817) \end{gathered}$ | $\begin{gathered} 0.8267^{* * *} \\ (0.2799) \end{gathered}$ |
| mid numeracy: female | $\begin{gathered} 0.9825^{* * *} \\ (0.3068) \end{gathered}$ | $\begin{gathered} 0.9888^{* * *} \\ (0.3049) \end{gathered}$ |
| high numeracy: female | $\begin{gathered} 0.8279^{* * *} \\ (0.2684) \end{gathered}$ | $\begin{gathered} 0.8352^{* * *} \\ (0.2661) \end{gathered}$ |
| age: male | $\begin{gathered} 0.0855 \\ (0.1346) \end{gathered}$ | $\begin{gathered} 0.0795 \\ (0.1354) \end{gathered}$ |
| age ${ }^{2}$ : male | $\begin{gathered} -0.0005 \\ (0.0011) \end{gathered}$ | $\begin{aligned} & -0.0004 \\ & (0.0011) \end{aligned}$ |
| age diff $>10$ | $\begin{gathered} 0.0453 \\ (0.2076) \end{gathered}$ | $\begin{gathered} 0.0317 \\ (0.2077) \end{gathered}$ |
| $2^{\text {nd }}$ income quartile | $\begin{gathered} 0.1014 \\ (0.1693) \end{gathered}$ | $\begin{gathered} 0.1069 \\ (0.1686) \end{gathered}$ |
| $3^{r d}$ income quartile | $\begin{gathered} 0.2703 \\ (0.1731) \end{gathered}$ | $\begin{gathered} 0.2629 \\ (0.1719) \end{gathered}$ |
| $4^{t h}$ income quartile | $\begin{gathered} 0.2668 \\ (0.1833) \end{gathered}$ | $\begin{gathered} 0.2623 \\ (0.1831) \end{gathered}$ |
| $2^{n d}$ wealth | $\begin{gathered} 0.7760^{* * *} \\ (0.1616) \end{gathered}$ | $\begin{gathered} 0.7802^{* * *} \\ (0.1612) \end{gathered}$ |
| $3^{r d}$ wealth quartile | $\begin{gathered} 0.8108^{* * *} \\ (0.1596) \end{gathered}$ | $\begin{gathered} 0.8157^{* * *} \\ (0.1588) \end{gathered}$ |
| $4^{t h}$ wealth quartile | $\begin{gathered} 1.2758^{* * *} \\ (0.2040) \end{gathered}$ | $\begin{gathered} 1.2812^{* * *} \\ (0.2014) \end{gathered}$ |
| in work: male | $\begin{gathered} -0.1609 \\ (0.1439) \end{gathered}$ | $\begin{gathered} -0.1676 \\ (0.1440) \end{gathered}$ |
| in work: female | $\begin{aligned} & -0.0106 \\ & (0.1306) \end{aligned}$ | $\begin{array}{r} -0.0254 \\ (0.1298) \end{array}$ |
| mid education: male | $\begin{gathered} 0.1481 \\ (0.1487) \end{gathered}$ | $\begin{gathered} 0.1571 \\ (0.1485) \end{gathered}$ |
| high education: male | $\begin{gathered} 0.2378 \\ (0.1928) \end{gathered}$ | $\begin{gathered} 0.2387 \\ (0.1924) \end{gathered}$ |
| mid education: female | $\begin{gathered} 0.0388 \\ (0.1537) \end{gathered}$ | $\begin{gathered} 0.0379 \\ (0.1539) \end{gathered}$ |
| high education: female | $\begin{gathered} 0.1064 \\ (0.2060) \end{gathered}$ | $\begin{gathered} 0.1064 \\ (0.2056) \end{gathered}$ |
| sickness index: male | $\begin{gathered} -0.0194 \\ (0.0327) \end{gathered}$ | $\begin{gathered} -0.0170 \\ (0.0325) \end{gathered}$ |
| sickness index: female | $\begin{gathered} -0.0372 \\ (0.0286) \end{gathered}$ | $\begin{gathered} -0.0427 \\ (0.0284) \end{gathered}$ |
| financial risk: male | $\begin{gathered} 0.0104 \\ (0.0263) \end{gathered}$ |  |
| financial patience: male | $\begin{gathered} 0.0104 \\ (0.0288) \end{gathered}$ |  |
| financial risk: female | $\begin{gathered} -0.0087 \\ (0.0251) \end{gathered}$ |  |
| financial patience: female | $\begin{gathered} 0.0432 \\ (0.0272) \end{gathered}$ |  |
| financial risk: hh |  | $\begin{gathered} 0.0105 \\ (0.0296) \end{gathered}$ |
| financial patience: hh |  | $\begin{gathered} 0.0265 \\ (0.0339) \end{gathered}$ |
| Constant | $\begin{array}{r} -5.4741 \\ (4.1672) \\ \hline \end{array}$ | $\begin{array}{r} -5.1140 \\ (4.1851) \\ \hline \end{array}$ |
| Number of observation | 688 | 688 |

Table 3: Heckman selection equation: household probability of holding risky assets. Baseline model selection: financial and general risk tolerance. Partners characteristics by gender.

| Under 65 years |  |  |
| :---: | :---: | :---: |
| share of risky assets | Financial risk measure (1) | General risk measure (2) |
| age: male | $\begin{gathered} -0.0876^{*} \\ (0.0469) \end{gathered}$ | $\begin{gathered} -0.0904^{*} \\ (0.0470) \end{gathered}$ |
| age $^{2}$ : male | $\begin{aligned} & 0.0007^{*} \\ & (0.0004) \end{aligned}$ | $\begin{gathered} 0.0007^{* *} \\ (0.0004) \end{gathered}$ |
| age diff $>10$ | $\begin{aligned} & -0.0580 \\ & (0.0531) \end{aligned}$ | $\begin{aligned} & -0.0741 \\ & (0.0532) \end{aligned}$ |
| $2^{\text {nd }}$ income quartile | $\begin{gathered} 0.0590 \\ (0.0463) \end{gathered}$ | $\begin{gathered} 0.0586 \\ (0.0463) \end{gathered}$ |
| $3^{r d}$ income quartile | $\begin{gathered} 0.0008 \\ (0.0479) \end{gathered}$ | $\begin{aligned} & -0.0006 \\ & (0.0473) \end{aligned}$ |
| $4^{\text {th }}$ income quartile | $\begin{gathered} 0.0243 \\ (0.0484) \end{gathered}$ | $\begin{gathered} 0.0254 \\ (0.0477) \end{gathered}$ |
| $2^{n d}$ wealth quartile | $\begin{gathered} 0.0810 \\ (0.0800) \end{gathered}$ | $\begin{gathered} 0.0773 \\ (0.0804) \end{gathered}$ |
| $3^{r d}$ wealth quartile | $\begin{aligned} & 0.1532^{*} \\ & (0.0805) \end{aligned}$ | $\begin{aligned} & 0.1566^{*} \\ & (0.0811) \end{aligned}$ |
| $4^{t h}$ wealth quartile | $\begin{gathered} 0.2027^{* *} \\ (0.0999) \end{gathered}$ | $\begin{aligned} & 0.1956^{*} \\ & (0.1008) \end{aligned}$ |
| in work: male | $\begin{aligned} & -0.0426 \\ & (0.0353) \end{aligned}$ | $\begin{aligned} & -0.0503 \\ & (0.0351) \end{aligned}$ |
| in work: female | $\begin{gathered} -0.0050 \\ (0.0315) \end{gathered}$ | $\begin{gathered} 0.0036 \\ (0.0314) \end{gathered}$ |
| mid education: male | $\begin{gathered} 0.0220 \\ (0.0413) \end{gathered}$ | $\begin{gathered} 0.0228 \\ (0.0416) \end{gathered}$ |
| high education: male | $\begin{gathered} 0.0163 \\ (0.0494) \end{gathered}$ | $\begin{gathered} 0.0150 \\ (0.0495) \end{gathered}$ |
| mid education: female | $\begin{gathered} -0.0454 \\ (0.0423) \end{gathered}$ | $\begin{gathered} -0.0574 \\ (0.0425) \end{gathered}$ |
| high education: female | $\begin{gathered} 0.0749 \\ (0.0499) \end{gathered}$ | $\begin{gathered} 0.0662 \\ (0.0501) \end{gathered}$ |
| sickness index: male | $\begin{gathered} 0.0149 \\ (0.0095) \end{gathered}$ | $\begin{aligned} & 0.0157^{*} \\ & (0.0095) \end{aligned}$ |
| sickness index: female | $\begin{aligned} & -0.0028 \\ & (0.0087) \end{aligned}$ | $\begin{gathered} -0.0027 \\ (0.0088) \end{gathered}$ |
| financial risk: male | $\begin{gathered} 0.0201^{* * *} \\ (0.0067) \end{gathered}$ |  |
| financial patience: male | $\begin{aligned} & -0.0043 \\ & (0.0072) \end{aligned}$ |  |
| financial risk: female | $\begin{gathered} 0.0186^{* * *} \\ (0.0064) \end{gathered}$ |  |
| financial patience: female | $\begin{gathered} 0.0039 \\ (0.0072) \end{gathered}$ |  |
| financial risk: hh |  | $\begin{gathered} 0.0333^{* * *} \\ (0.0073) \end{gathered}$ |
| financial patience: hh |  | $\begin{gathered} -0.0022 \\ (0.0087) \end{gathered}$ |
| Constant | $\begin{aligned} & 2.7789^{*} \\ & (1.5173) \\ & \hline \end{aligned}$ | $\begin{aligned} & 2.9300^{*} \\ & (1.5183) \\ & \hline \end{aligned}$ |
| Inverse Mills Ratio $\lambda$ | $\begin{gathered} 0.0384 \\ (0.1467) \\ \hline \end{gathered}$ | $\begin{gathered} 0.0250 \\ (0.1462) \\ \hline \end{gathered}$ |
| Number of observation | 688 | 688 |

## C. 2 Male and female - exclusion restriction

The Heckman methodology shows to be sensible to the exclusion restriction. This Section provides evidence that the results are robust even when using different exclusion restriction: in particular, here only the numeracy of the husband plays the role of the exclusion restriction. The results are consistent (even if sometimes attenuated) and the Invers Mills Ratio become significant in one case, confirming the relevance of controlling for incidentally truncated variables bias.

Table 4: Household portfolio determinants - collective approach. Heckman selection equation: household probability of holding risky assets. Partners characteristics by gender.

| participation | Individual risk <br> (1) | $\begin{aligned} & \hline \hline \text { Weighted risk } \\ & (2) \end{aligned}$ |
| :---: | :---: | :---: |
| low numeracy: male | 0.2240 | $0.5016^{* *}$ |
|  | (0.3885) | (0.2227) |
| mid numeracy: male | 0.4221 | $0.6777^{* * *}$ |
|  | (0.4005) | (0.2329) |
| high numeracy: male | 0.5185 | 0.7059 *** |
|  | (0.3684) | (0.2092) |
| age: male | 0.0779 | $0.1279^{* *}$ |
|  | (0.1348) | (0.0633) |
| age ${ }^{2}$ : male | -0.0004 | -0.0009* |
|  | (0.0011) | (0.0005) |
| age diff $>10$ | 0.0424 | 0.1065 |
|  | (0.2040) | (0.1649) |
| $2^{n d}$ income quartile | 0.1313 | $0.2381{ }^{* *}$ |
|  | (0.1667) | (0.1058) |
| $3^{r d}$ income quartile | 0.2727 | $0.3451^{* * *}$ |
|  | (0.1711) | (0.1171) |
| $4^{t h}$ income quartile | 0.2784 | $0.3567^{* * *}$ |
|  | (0.1819) | (0.1337) |
| $2^{n d}$ wealth quartile | $0.7739^{* * *}$ | $0.7516^{* * *}$ |
|  | (0.1600) | (0.1065) |
| $3^{r d}$ wealth quartile | $0.8059^{* * *}$ | $0.9090^{* * *}$ |
|  | (0.1578) | (0.1135) |
| $4^{t h}$ wealth quartile | $1.2966^{* * *}$ | $1.4444^{* * *}$ |
|  | (0.2024) | (0.1432) |
| in work: male | -0.1389 | -0.2144* |
|  | (0.1423) | (0.1118) |
| in work: female | 0.0092 | -0.0105 |
|  | (0.1296) | (0.1075) |
| mid education: male | 0.1420 | $0.2856{ }^{* * *}$ |
|  | (0.1468) | (0.0967) |
| high education: male | 0.2235 | $0.3110^{* *}$ |
|  | (0.1906) | (0.1359) |
| mid education: female | $0.0380$ | $0.0656$ |
|  | (0.1516) | $(0.0945)$ |
| high education: female | 0.1267 | 0.1585 |
|  | (0.2029) | (0.1439) |
| sickness index: male | -0.0113 | -0.0178 |
|  | (0.0322) | (0.0190) |
| sickness index: female | -0.0519* | -0.0214 |
|  | (0.0276) | (0.0177) |
| financial risk: male | 0.0180 |  |
|  | (0.0258) |  |
| financial patience: male | $0.0141$ |  |
|  | (0.0284) |  |
| financial risk: female | -0.0191 |  |
|  | (0.0246) |  |
| financial patience: female | 0.0375 |  |
|  | (0.0269) |  |
| financial risk: hh |  | 0.0162 |
|  |  | (0.0204) |
| financial patience: hh |  | -0.0203 |
|  |  | (0.0235) |
| Constant | -4.4250 | $-5.7125^{* *}$ |
|  | (4.1657) | (2.2274) |
| Number of observations | 1,442 | 1,442 |

Table 5: Household portfolio determinants - collective approach. Heckman outcome equation: household probability of holding risky assets. Partners characteristics by gender.

| share of risky assets | Individual risk <br> (1) | Weighted risk <br> (2) |
| :---: | :---: | :---: |
| age: male | $\begin{gathered} -0.0835^{*} \\ (0.0477) \end{gathered}$ | $\begin{gathered} 0.0234 \\ (0.0206) \end{gathered}$ |
| age ${ }^{2}$ : male | $\begin{aligned} & 0.0007^{*} \\ & (0.0004) \end{aligned}$ | $\begin{gathered} -0.0001 \\ (0.0001) \end{gathered}$ |
| age diff $>10$ | $\begin{aligned} & -0.0558 \\ & (0.0540) \end{aligned}$ | $\begin{gathered} 0.0201 \\ (0.0433) \end{gathered}$ |
| $2^{\text {nd }}$ income quartile | $\begin{gathered} 0.0649 \\ (0.0482) \end{gathered}$ | $\begin{gathered} 0.0248 \\ (0.0342) \end{gathered}$ |
| $3^{r d}$ income quartile | $\begin{gathered} 0.0128 \\ (0.0536) \end{gathered}$ | $\begin{gathered} 0.0379 \\ (0.0373) \end{gathered}$ |
| $4^{t h}$ income quartile | $\begin{gathered} 0.0366 \\ (0.0545) \end{gathered}$ | $\begin{gathered} 0.0227 \\ (0.0388) \end{gathered}$ |
| $2^{\text {nd }}$ wealth quartile | $\begin{gathered} 0.1217 \\ (0.1115) \end{gathered}$ | $\begin{gathered} 0.1416^{* *} \\ (0.0703) \end{gathered}$ |
| $3^{\text {rd }}$ wealth quartile | $\begin{aligned} & 0.1948^{*} \\ & (0.1134) \end{aligned}$ | $\begin{gathered} 0.2575^{* * *} \\ (0.0792) \end{gathered}$ |
| $4^{\text {th }}$ wealth quartile | $\begin{aligned} & 0.2575^{*} \\ & (0.1447) \end{aligned}$ | $\begin{gathered} 0.3451^{* * *} \\ (0.0983) \end{gathered}$ |
| in work: male | $\begin{gathered} -0.0474 \\ (0.0369) \end{gathered}$ | $\begin{aligned} & -0.0500^{*} \\ & (0.0300) \end{aligned}$ |
| in work: female | $\begin{gathered} -0.0041 \\ (0.0321) \end{gathered}$ | $\begin{gathered} -0.0019 \\ (0.0270) \end{gathered}$ |
| mid education: male | $\begin{gathered} 0.0294 \\ (0.0443) \end{gathered}$ | $\begin{gathered} 0.0493 \\ (0.0319) \end{gathered}$ |
| high education: male | $\begin{gathered} 0.0255 \\ (0.0533) \end{gathered}$ | $\begin{gathered} 0.0451 \\ (0.0378) \end{gathered}$ |
| mid education: female | $\begin{gathered} -0.0434 \\ (0.0430) \end{gathered}$ | $\begin{gathered} 0.0328 \\ (0.0262) \end{gathered}$ |
| high education. female | $\begin{gathered} 0.0793 \\ (0.0513) \end{gathered}$ | ${ }_{(0.0344)}^{0.1117^{* * *}}$ |
| sickness index. male | $\begin{gathered} 0.0143 \\ (0.0097) \end{gathered}$ | $\begin{aligned} & -0.0016 \\ & (0.0058) \end{aligned}$ |
| sickness index: female | $\begin{aligned} & -0.0055 \\ & (0.0103) \end{aligned}$ | $\begin{aligned} & -0.0043 \\ & (0.0052) \end{aligned}$ |
| financial risk: male | $\begin{gathered} 0.0209^{* * *} \\ (0.0069) \end{gathered}$ |  |
| financial patience: male | $\begin{aligned} & -0.0037 \\ & (0.0075) \end{aligned}$ |  |
| financial risk: female | $\begin{gathered} 0.0178^{* * *} \\ (0.0066) \end{gathered}$ |  |
| financial patience: female | $\begin{gathered} 0.0050 \\ (0.0076) \end{gathered}$ |  |
| financial risk: hh |  | $\begin{gathered} 0.0316^{* * *} \\ (0.0054) \end{gathered}$ |
| financial patience: hh |  | $\begin{aligned} & -0.0011 \\ & (0.0063) \end{aligned}$ |
| Constant | $\begin{gathered} 2.5169 \\ (1.5973) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.9818 \\ (0.7989) \\ \hline \end{array}$ |
| ${ }_{\lambda}^{\text {Inverse Mills Ratio }}$ | $\begin{gathered} 0.1277 \\ (0.2232) \\ \hline \end{gathered}$ | $\begin{aligned} & 0.2365^{*} \\ & (0.1398) \end{aligned}$ |
| Number of observations | 1,442 | 1,442 |

## C. 3 General risk preferences

Table 6: Household portfolio determinants - collective approach. Heckman selection equation: household probability of holding risky assets. Partners characteristics by gender.

| participation | Individual risk <br> (1) | Weighted risk (2) |
| :---: | :---: | :---: |
| low numeracy: male | $0.4519^{* *}$ | $0.4569^{* *}$ |
|  | (0.2249) | (0.2247) |
| mid numeracy: male | $0.6237 * * *$ | 0.6342*** |
|  | (0.2353) | (0.2351) |
| high numeracy: male | $0.6543^{* * *}$ | $0.6545^{* * *}$ |
|  | (0.2118) | (0.2116) |
| low numeracy: female | $0.4098 * *$ | $0.4017^{* *}$ |
|  | (0.1803) | (0.1802) |
| mid numeracy: female | $0.3810^{* *}$ | $0.3822^{* *}$ |
|  | $(0.1942)$ | (0.1942) |
| high numeracy: female | 0.4088** | $0.4079^{* *}$ |
|  | (0.1712) | (0.1712) |
| age: male | 0.1281** | $0.1274^{* *}$ |
|  | (0.0632) | (0.0634) |
| age ${ }^{2}$ : male | -0.0009* | -0.0009* |
|  | (0.0005) | (0.0005) |
| age diff $>10$ | 0.1208 | 0.1056 |
|  | (0.1659) | (0.1660) |
| $2^{n d}$ income quartile | 0.2392** | 0.2396** |
|  | (0.1065) | (0.1063) |
| $3^{\text {rd }}$ income quartile | $0.3556 * * *$ | $0.3466^{* * *}$ |
|  | (0.1177) | (0.1174) |
| $4^{t h}$ income quartile | $0.3508^{* * *}$ | 0.3579 *** |
|  | (0.1343) | (0.1340) |
| $2^{n d}$ wealth quartile | $0.7364^{* * *}$ | $0.7363^{* * *}$ |
|  | (0.1069) | (0.1069) |
| $3^{\text {rd }}$ wealth quartile | $0.8943^{* * *}$ | 0.8960 *** |
|  | (0.1142) | (0.1139) |
| $4^{\text {th }}$ wealth quartile | 1.4250 *** | $1.4210^{* * *}$ |
|  | (0.1448) | (0.1437) |
| in work: male | -0.2224** | -0.2302** |
|  | (0.1123) | (0.1122) |
| in work: female | -0.0120 | -0.0193 |
|  | (0.1084) | (0.1078) |
| mid education: male | $0.3024^{* * *}$ | $0.3008^{* * *}$ |
|  | (0.0974) | (0.0973) |
| high education: male | 0.3383** | $0.3254^{* *}$ |
|  | (0.1372) | (0.1368) |
| mid education: female | 0.0566 | 0.0537 |
|  | (0.0951) | (0.0950) |
| high education: female | 0.1433 | 0.1471 |
|  | (0.1454) | (0.1451) |
| sickness index: male | -0.0201 | -0.0192 |
|  | (0.0191) | (0.0190) |
| sickness index: female | -0.0128 | -0.0157 |
|  | (0.0182) | (0.0181) |
| general risk: male | 0.0177 |  |
|  | (0.0172) |  |
| general patience: male | -0.0281 |  |
|  | (0.0196) |  |
| general risk: female | -0.0052 |  |
|  | (0.0176) |  |
| general patience: female | 0.0260 |  |
|  | (0.0183) |  |
| general risk: hh |  | $0.0173$ |
|  |  | (0.0205) |
| general patience: hh |  | -0.0187 |
|  |  | (0.0236) |
| Constant | $-6.1690^{* * *}$ | $\begin{gathered} -5.9974^{* * *} \\ (2.2331) \end{gathered}$ |
|  | (2.2245) | (2.2331) |
| Number of observations | 1,442 | 1,442 |

Table 7: Household portfolio determinants - collective approach. Heckman outcome equation: household probability of holding risky assets. Partners characteristics by gender.

| share of risky asset | Individual risk <br> (1) | $\begin{aligned} & \hline \text { Weighted risk } \\ & \text { (2) } \end{aligned}$ |
| :---: | :---: | :---: |
| age: male | 0.0147 | 0.0143 |
|  | (0.0197) | (0.0197) |
| age ${ }^{2}$ : male | -0.0001 | -0.0001 |
|  | (0.0001) | (0.0001) |
| age diff $>10$ | 0.0224 | 0.0145 |
|  | (0.0413) | (0.0413) |
| $2^{n d}$ income quartile | 0.0105 | 0.0075 |
|  | (0.0321) | (0.0322) |
| $3^{r d}$ income quartile | 0.0165 | 0.0146 |
|  | (0.0350) | (0.0349) |
| $4^{t h}$ income quartile | 0.0028 | -0.0001 |
|  | (0.0359) | (0.0362) |
| $2^{n d}$ wealth quartile | 0.0744 | 0.0752 |
|  | (0.0650) | (0.0657) |
| $3^{r d}$ wealth quartile | 0.1801** | 0.1807** |
|  | (0.0729) | (0.0738) |
| $4^{t h}$ wealth quartile | 0.2487*** | $0.2464^{* * *}$ |
|  | (0.0900) | (0.0910) |
| in work: male | -0.0329 | -0.0390 |
|  | (0.0283) | (0.0284) |
| in work: male | -0.0111 | -0.0016 |
|  | (0.0257) | (0.0256) |
| mid education: male | 0.0298 | 0.0298 |
|  | (0.0301) | (0.0301) |
| high education: male | 0.0267 | 0.0249 |
|  | (0.0358) | (0.0356) |
| mid education: female | 0.0302 | 0.0282 |
|  | (0.0250) | (0.0250) |
| high education: female | $0.1005^{* * *}$ | $0.1041^{* * *}$ |
|  | (0.0324) | (0.0324) |
| sickness index: male | -0.0002 | 0.0002 |
|  | (0.0056) | (0.0056) |
| sickness index: female | -0.0032 | -0.0027 |
|  | (0.0049) | (0.0049) |
| general risk: male | $0.0158^{* * *}$ |  |
|  | (0.0044) |  |
| general patience: male | -0.0018 |  |
|  | (0.0049) |  |
| general risk: female | $0.0187^{* * *}$ |  |
|  | $(0.0045)$ |  |
| general patience: female | 0.0016 |  |
|  | (0.0048) |  |
| general risk: hh |  | $0.0307^{* * *}$ |
|  |  | (0.0051) |
| general patience: hh |  | -0.0002 |
|  |  | (0.0060) |
| Constant | -0.5320 | -0.4960 |
|  | (0.7598) | (0.7603) |
| Inverse Mills Ratio |  |  |
| $\lambda$ | 0.0876 | 0.0849 |
|  | (0.1276) | (0.1290) |
| Number of observations | 1,442 | 1,442 |

## C. 4 Partners characteristics by financial respondent

In this Section, I identify the (unitary) household risk preferences with the one of the household head, where the household head is the financial respondent of the survey. Moreover, I distinguish the partners demographics by financial and non-financial respondent.

## C.4.1 All sample

Table 8: Household portfolio determinants - collective approach. Heckman outcome equation: household probability of holding risky assets. Partners characteristics by financial respondent.

| participation | Individual risk <br> (1) | $\underset{\text { (2) }}{\text { Weighted risk }}$ |
| :---: | :---: | :---: |
| low numeracy: respondent | $\begin{gathered} 0.5381^{* *} \\ (0.2535) \end{gathered}$ | $\begin{gathered} 0.5464^{* *} \\ (0.2529) \end{gathered}$ |
| mid numeracy: respondent | $0.5567^{* *}$ | $0.5623^{* *}$ |
|  | (0.2646) | (0.2640) |
| high numeracy: respondent | $0.5856^{* *}$ | $0.5923^{* *}$ |
| low numeracy: non-respondent | 0.3739** | $0.3675^{* *}$ |
|  | (0.1716) | (0.1711) |
| mid numeracy: non-respondent | $0.4324^{* *}$ | $0.4329^{* *}$ |
|  | (0.1830)* | (0.1826) |
| high numeracy: non-respondent | $0.4843^{* * *}$ | $0.4845^{* * *}$ |
| age: respondent | 0.0477 | 0.0468 |
|  | (0.0610) | (0.0608) |
| age ${ }^{2}$ : respondent | $\begin{gathered} -0.0004 \\ (0.0004) \end{gathered}$ | $\begin{gathered} -0.0003 \\ (0.0004) \end{gathered}$ |
| age difference > 10 | 0.0547 | 0.0546 |
|  | (0.1634) | (0.1634) |
| $2^{\text {nd }}$ hh income quartile | $0.2573^{* *}$ | $0.2563 * *$ |
|  | (0.1064) | (0.1061) |
| $3^{\text {rd }}$ hh income quartile | $0.3895^{* * *}$ | $0.3896^{* * *}$ |
| $4^{t h} \mathrm{hh}$ income quartile | 0.3914*** | 0.3933*** |
|  | (0.1335) | (0.1335) |
| $2^{\text {nd }}$ hh wealth quartile | $0.7172^{* * *}$ | $0.7258^{* * *}$ |
|  | (0.1068) | (0.1067) |
| $3^{\text {rd }}$ hh wealth quartile | $0.9035 * * *$ | $0.9099^{* * *}$ |
|  | (0.1151) | (0.1149) |
| $4^{\text {th }}$ hh wealth quartile | 1.4201*** | 1.4290*** |
|  | (0.1452) | (0.1451) |
| in work: male: respondent | -0.1655 | -0.1678 |
|  | (0.1094) | (0.1095) |
| in work: male: non-respondent | -0.2034* | -0.2077* |
|  | (0.1079) | (0.1077) |
| mid education: respondent | 0.2719*** | 0.2694*** |
|  | (0.0965) | (0.0965) |
| high education: respondent | $0.3466{ }^{* *}$ | $0.3418^{* *}$ |
|  | (0.1376) | (0.1375) |
| mid education: non-respondent | $0.0514$ | $0.0501$ |
| high education: non-respondent | 0.1159 | 0.1209 |
|  | (0.1457) | (0.1455) |
| sickness: respondent | -0.0389* | -0.0387* |
|  | (0.0205) | (0.0204) |
| sickness: non-respondent | -0.0010 | -0.0017 |
| female financial respondent | $(0.0173)$ 0.0905 | $(0.0173)$ 0.0811 |
|  | (0.0880) | (0.0867) |
| financial risk: respondent | $0.0168$ |  |
| financial risk: non-respondent | 0.0018 |  |
|  | (0.0175) |  |
| financial patience: respondent | -0.0072 |  |
|  | (0.0193) |  |
| financial patience: non-respondent | 0.0099 |  |
|  | (0.0183) |  |
| financial risk: hh |  | 0.0198 |
|  |  | (0.0205) |
| financial patience: hh |  | $-0.0141$ |
| Constant | -3.1486 | -2.9982 |
|  | (2.1313) | (2.1256) |
| Number of observations | 1,442 | 1,442 |

Table 9: Household portfolio determinants - collective approach. Heckman outcome equation: household probability of holding risky assets. Partners characteristics by financial respondent.

| share of risky assets | Individual risk <br> (1) | $\underset{(2)}{\text { Weighted risk }}$ |
| :---: | :---: | :---: |
| age: respondent | -0.0061 | -0.0059 |
| age $^{2}$ : respondent | $(0.0163)$ 0.0001 | $(0.0163)$ 0.0001 |
|  | (0.0001) | (0.0001) |
| age difference $>10$ | 0.0266 | 0.0218 |
|  | (0.0409) | (0.0409) |
| $2^{\text {nd }} \mathrm{hh}$ income quartile | -0.0041 | -0.0052 |
|  | (0.0328) | (0.0327) |
| $3^{\text {rd }} \mathrm{hh}$ income quartile | $-0.0046$ | -0.0027 |
|  | (0.0364) | (0.0363) |
| $4^{\text {th }} \mathrm{hh}$ income quartile | -0.0184 | -0.0183 |
| $2^{\text {nd }} \mathrm{hh}$ wealth quartile |  |  |
| 2 hh wealth quartile | (0.0681) | (0.0685) |
| $3^{\text {rd }}$ hh wealth quartile | 0.1105 | $0.1301 *$ |
|  | (0.0780) | (0.0784) |
| $4^{\text {th }}$ hh wealth quartile | $0.1594 *$ | $0.1809 *$ |
|  | (0.0962) | (0.0967) |
| in work: male: respondent | -0.0061 | -0.0067 |
|  | (0.0269) | (0.0269) |
| in work: male: non-respondent | $\begin{aligned} & -0.0380 \\ & (0.0271) \end{aligned}$ | $\begin{aligned} & -0.0372 \\ & (0.0272) \end{aligned}$ |
| mid education: respondent | 0.0065 | 0.0092 |
|  | (0.0309) | (0.0308) |
| high education: respondent | $\begin{gathered} 0.0403 \\ (0.0365) \end{gathered}$ | $\begin{aligned} & 0.0452 \\ & (00363) \end{aligned}$ |
| mid education: non-respondent | 0.0270 | 0.0284 |
|  | (0.0245) | (0.0244) |
| high education: non-respondent | $\begin{gathered} 0.0521 \\ (0.0320) \end{gathered}$ | $0.0557^{*}$ |
| sickness: respondent | -0.0030 | -0.0034 |
|  | (0.0062) | (0.0062) |
| sickness: non-respondent | $\begin{gathered} 0.0012 \\ (0.0046) \end{gathered}$ | $\begin{aligned} & 0.0016 \\ & (0.0046) \end{aligned}$ |
| female financial respondent | -0.0087 | -0.0096 |
|  | (0.0218) | (0.0210) |
| financial risk: respondent | $\begin{gathered} 0.0197^{* * * *} \\ (0.0043) \end{gathered}$ |  |
| financial risk: non-respondent | $0.0139^{* * *}$ |  |
|  | (0.0044) |  |
| financial patience: respondent | $\begin{gathered} -0.0001 \\ (0.0047) \end{gathered}$ |  |
| financial patience: non-respondent | -0.0010 |  |
|  | (0.0047) |  |
| financial risk: hh |  | $0.0294{ }^{* * *}$ |
| financial patience: hh |  | -0.0006 |
|  |  | (0.0059) |
| Constant | 0.3664 | 0.3473 |
|  | (0.6045) | (0.6012) |
| Inverse Mills ratio |  |  |
| $\lambda$ | -0.0449 | -0.0168 |
|  | (0.1356) | (0.1359) |
| Number of observations | 1,442 | 1,442 |
| Selection | 1,070 | 1,070 |

Table 10: Collective vs unitary approach - likelihood ratio test. Heckman outcome equation: household probability of holding risky assets. Partners characteristics by financial respondent.

| share of risky assets | Collective models |  |
| :--- | :---: | :---: | :---: |
|  |  |  |
| (1) |  |  |\(\left.\quad \begin{array}{c}Unitary <br>

Household risk <br>
(2)\end{array}\right)\)

[^14]Table 11: Collective vs unitary approach - information criteria. Heckman outcome equation: household probability of holding risky assets. Partners characteristics by financial respondent.

| share of risky assets | Collective model <br> (1) | Unitary model <br> (2) |
| :---: | :---: | :---: |
| demographics $\square^{4}$ | * | * |
| financial risk: hh | $\begin{gathered} 0.0293^{* * *} \\ (0.0052) \end{gathered}$ |  |
| financial risk: respondent |  | $\begin{gathered} 0.0219^{* * *} \\ (0.0043) \end{gathered}$ |
| Constant | $\begin{gathered} 0.3705 \\ (0.6069) \end{gathered}$ | $\begin{gathered} 0.3381 \\ (0.5897) \end{gathered}$ |
| AIC | 1926 | 1933 |
| BIC | 2190 | 2197 |
| Number of observations | 1,442 | 1,442 |

[^15]
## C.4.2 Under 65 years

Table 12: Under 65 years: household portfolio determinants - collective approach. Heckman selection equation: household probability of holding risky assets. Partners characteristics by financial respondent.

| participation | $\begin{gathered} \hline \hline \text { Financial } \\ \text { risk measure } \\ (1) \end{gathered}$ | $\begin{aligned} & \hline \text { General } \\ & \text { risk measure } \\ & (2) \end{aligned}$ |
| :---: | :---: | :---: |
| low numeracy: respondent | $\begin{gathered} 1.5660^{* * *} \\ (0.5447) \end{gathered}$ | $\begin{gathered} 1.5861^{* * *} \\ (0.5372) \end{gathered}$ |
| mid numeracy: respondent | $\begin{gathered} 1.5498^{* * *} \\ (0.5577) \end{gathered}$ | $\begin{gathered} 1.5842^{* * *} \\ (0.5501) \end{gathered}$ |
| high numeracy: respondent | $\begin{gathered} 1.5636^{* * *} \\ (0.5329) \end{gathered}$ | $\begin{gathered} 1.5766^{* * *} \\ (0.5257) \end{gathered}$ |
| low numeracy: non-respondent | $\begin{aligned} & 0.4851^{*} \\ & (0.2609) \end{aligned}$ | $\begin{aligned} & 0.4776^{*} \\ & (0.2601) \end{aligned}$ |
| mid numeracy: non-respondent | $\begin{gathered} 0.6216^{* *} \\ (0.2844) \end{gathered}$ | $\begin{gathered} 0.6308^{* *} \\ (0.2834) \end{gathered}$ |
| high numeracy: non-respondent | $\begin{gathered} 0.6550^{* * *} \\ (0.2467) \end{gathered}$ | $\begin{gathered} 0.6633^{* * *} \\ (0.2453) \end{gathered}$ |
| age: respondent | $\begin{gathered} -0.0658 \\ (0.1523) \end{gathered}$ | $\begin{gathered} -0.0630 \\ (0.1516) \end{gathered}$ |
| age ${ }^{2}$ : respondent | $\begin{gathered} 0.0006 \\ (0.0013) \end{gathered}$ | $\begin{gathered} 0.0006 \\ (0.0013) \end{gathered}$ |
| age difference $>10$ | $\begin{gathered} -0.0102 \\ (0.2083) \end{gathered}$ | $\begin{gathered} -0.0034 \\ (0.2083) \end{gathered}$ |
| $2^{n d}$ income quartile | $\begin{gathered} 0.0844 \\ (0.1716) \end{gathered}$ | $\begin{gathered} 0.0845 \\ (0.1708) \end{gathered}$ |
| $3^{r d}$ income quartile | $\begin{gathered} 0.2582 \\ (0.1746) \end{gathered}$ | $\begin{gathered} 0.2616 \\ (0.1742) \end{gathered}$ |
| $4^{t h}$ income quartile | $\begin{gathered} 0.2337 \\ (0.1843) \end{gathered}$ | $\begin{gathered} 0.2348 \\ (0.1842) \end{gathered}$ |
| $2^{n d}$ wealth quartile | $\begin{gathered} 0.8295^{* * *} \\ (0.1614) \end{gathered}$ | $\begin{gathered} 0.8288^{* * *} \\ (0.1609) \end{gathered}$ |
| $3^{\text {rd }}$ wealth quartile | $\begin{gathered} 0.8495^{* * *} \\ (0.1612) \end{gathered}$ | $\begin{gathered} 0.8573^{* * *} \\ (0.1603) \end{gathered}$ |
| $4^{t h}$ wealth quartile | $\begin{gathered} 1.3504^{* * *} \\ (0.2058) \end{gathered}$ | $\begin{gathered} 1.3612^{* * *} \\ (0.2051) \end{gathered}$ |
| in work: male: respondent | $\begin{gathered} -0.0938 \\ (0.1378) \end{gathered}$ | $\begin{gathered} -0.0976 \\ (0.1376) \end{gathered}$ |
| in work: male: non-respondent | $\begin{gathered} -0.1488 \\ (0.1328) \end{gathered}$ | $\begin{gathered} -0.1611 \\ (0.1322) \end{gathered}$ |
| mid education: respondent | $\begin{gathered} 0.2233 \\ (0.1534) \end{gathered}$ | $\begin{gathered} 0.2355 \\ (0.1530) \end{gathered}$ |
| high education: respondent | $\begin{aligned} & 0.3714^{*} \\ & (0.1984) \end{aligned}$ | $\begin{aligned} & 0.3691^{*} \\ & (0.1980) \end{aligned}$ |
| mid education: non-respondent | $\begin{gathered} -0.1326 \\ (0.1536) \end{gathered}$ | $\begin{gathered} -0.1169 \\ (0.1523) \end{gathered}$ |
| high education: non-respondent | $\begin{gathered} -0.0951 \\ (0.2075) \end{gathered}$ | $\begin{gathered} -0.0711 \\ (0.2056) \end{gathered}$ |
| sickness index: respondent | $\begin{gathered} -0.0707^{* *} \\ (0.0327) \end{gathered}$ | $\begin{gathered} -0.0695^{* *} \\ (0.0323) \end{gathered}$ |
| sickness index: non-respondent | $\begin{gathered} -0.0011 \\ (0.0284) \end{gathered}$ | $\begin{gathered} -0.0040 \\ (0.0283) \end{gathered}$ |
| female respondent | $\begin{gathered} 0.1558 \\ (0.1301) \end{gathered}$ | $\begin{gathered} 0.1315 \\ (0.1274) \end{gathered}$ |
| financial risk: respondent | $\begin{gathered} 0.0172 \\ (0.0259) \end{gathered}$ |  |
| financial risk: non-respondent | $\begin{gathered} -0.0061 \\ (0.0253) \end{gathered}$ |  |
| financial patience: respondent | $\begin{gathered} 0.0054 \\ (0.0288) \end{gathered}$ |  |
| financial patience: non-respondent | $\begin{aligned} & 0.0500^{*} \\ & (0.0274) \end{aligned}$ |  |
| financial risk: hh |  | $\begin{gathered} 0.0158 \\ (0.0296) \end{gathered}$ |
| financial patience: hh |  | $\begin{gathered} 0.0313 \\ (0.0337) \end{gathered}$ |
| Constant | $\begin{aligned} & -1.0774 \\ & (4.6057) \\ & \hline \end{aligned}$ | $\begin{aligned} & -1.0133 \\ & (4.5937) \\ & \hline \end{aligned}$ |
| Number of observations | 688 | 688 |

Table 13: Under 65 years: household portfolio determinants - collective approach. Heckman outcome equation: household probability of holding risky assets. Partners characteristics by financial respondent.

| participation | $\begin{aligned} & \text { Financial } \\ & \text { risk measure } \\ & (1) \end{aligned}$ | General risk measure (2) |
| :---: | :---: | :---: |
| age: respondent | $\begin{aligned} & -0.0221 \\ & (0.0404) \end{aligned}$ | $\begin{gathered} -0.0189 \\ (0.0404) \end{gathered}$ |
| age $^{2}:$ respondent age difference $>10$ | $\begin{gathered} 0.0002 \\ (0.0003) \\ -0.0071 \\ (0.0510) \end{gathered}$ | $\begin{gathered} 0.0001 \\ (0.0003) \\ -0.0160 \\ (0.0513) \end{gathered}$ |
| $2^{\text {nd }}$ income quartile | $\begin{gathered} 0.0495 \\ (0.0464) \end{gathered}$ | $\begin{gathered} 0.0487 \\ (0.0463) \end{gathered}$ |
| $3^{\text {rd }}$ income quartile | $\begin{gathered} -0.0078 \\ (0.0477) \end{gathered}$ | $\begin{gathered} -0.0048 \\ (0.0477) \end{gathered}$ |
| $4^{\text {th }}$ income quartile | $\begin{gathered} 0.0074 \\ (0.0477) \end{gathered}$ | $\begin{gathered} 0.0109 \\ (0.0478) \end{gathered}$ |
| $2^{n d}$ wealth quartile | $\begin{gathered} 0.0499 \\ (0.0877) \end{gathered}$ | $\begin{gathered} 0.0500 \\ (0.0873) \end{gathered}$ |
| $3^{r d}$ wealth quartile | $\begin{aligned} & 0.1462^{*} \\ & (0.0881) \end{aligned}$ | $\begin{aligned} & 0.1527^{*} \\ & (0.0881) \end{aligned}$ |
| $4^{t h}$ wealth quartile | $\begin{gathered} 0.1770 \\ (0.1118) \end{gathered}$ | $\begin{gathered} 0.1775 \\ (0.1121) \end{gathered}$ |
| in work: male: respondent | $\begin{gathered} 0.0253 \\ (0.0334) \end{gathered}$ | $\begin{gathered} 0.0255 \\ (0.0335) \end{gathered}$ |
| in work: male: non-respondent | $\begin{gathered} -0.0750^{* *} \\ (0.0331) \end{gathered}$ | $\begin{gathered} -0.0737^{* *} \\ (0.0333) \end{gathered}$ |
| mid education: respondent | $\begin{gathered} 0.0132 \\ (0.0471) \end{gathered}$ | $\begin{gathered} 0.0022 \\ (0.0477) \end{gathered}$ |
| high education: respondent | $\begin{gathered} 0.0710 \\ (0.0554) \end{gathered}$ | $\begin{gathered} 0.0610 \\ (0.0557) \end{gathered}$ |
| mid education: non-respondent | $\begin{aligned} & -0.0234 \\ & (0.0405) \end{aligned}$ | $\begin{gathered} -0.0208 \\ (0.0405) \end{gathered}$ |
| high education: non-respondent | $\begin{gathered} 0.0167 \\ (0.0477) \end{gathered}$ | $\begin{gathered} 0.0213 \\ (0.0477) \end{gathered}$ |
| sickness index: respondent | $\begin{gathered} 0.0069 \\ (0.0112) \end{gathered}$ | $\begin{gathered} 0.0064 \\ (0.0111) \end{gathered}$ |
| sickness index: non-respondent | $\begin{gathered} 0.0071 \\ (0.0080) \end{gathered}$ | $\begin{gathered} 0.0082 \\ (0.0080) \end{gathered}$ |
| female respondent | $\begin{aligned} & -0.0263 \\ & (0.0331) \end{aligned}$ | $\begin{gathered} -0.0237 \\ (0.0318) \end{gathered}$ |
| financial risk: respondent | $\begin{gathered} 0.0191^{* * *} \\ (0.0063) \end{gathered}$ |  |
| financial risk: non-respondent | $\begin{gathered} 0.0176^{* * *} \\ (0.0065) \end{gathered}$ |  |
| financial patience: respondent | $\begin{aligned} & -0.0014 \\ & (0.0072) \end{aligned}$ |  |
| financial patience: non-respondent | $\begin{gathered} -0.0018 \\ (0.0077) \end{gathered}$ |  |
| financial risk: hh |  | $\begin{gathered} 0.0296^{* * *} \\ (0.0073) \end{gathered}$ |
| financial patience: hh |  | $\begin{gathered} -0.0051 \\ (0.0088) \end{gathered}$ |
| Constant | $\begin{gathered} 0.9267 \\ (1.2007) \\ \hline \end{gathered}$ | $\begin{gathered} 0.8873 \\ (1.2033) \\ \hline \end{gathered}$ |
| Inverse Mills Ratio $\lambda$ | $\begin{array}{r} -0.0096 \\ (0.1577) \\ \hline \end{array}$ | $\begin{array}{r} -0.0130 \\ (0.1572) \\ \hline \end{array}$ |
| Number of observations | 688 | 688 |

Table 14: Under 65 years: collective vs unitary approach - likelihood ratio test. Heckman outcome equation: household share of net financial wealth allocated in risky assets. Partners characteristics by financial respondent.

| Under 65 years |  |  |  |
| :---: | :---: | :---: | :---: |
|  | Collective models |  |  |
| share of risky assets | Partner risk <br> (1) | Household risk <br> (2) | Unodey <br> (3) |
| demographics $\int^{a}$ | * | * | * |
| financial risk: hh |  | $\begin{gathered} 0.0225^{* *} \\ (0.0111) \end{gathered}$ |  |
| financial risk: respondent | $\begin{gathered} 0.0196^{* * *} \\ (0.0066) \end{gathered}$ | $\begin{gathered} 0.0029 \\ (0.0104) \end{gathered}$ | $\begin{gathered} 0.0235^{* * *} \\ (0.0065) \end{gathered}$ |
| financial risk: non-respondent | $\begin{gathered} 0.0188^{* * *} \\ (0.0068) \end{gathered}$ |  |  |
| Constant | $\begin{gathered} 0.9647 \\ (1.2533) \\ \hline \end{gathered}$ | $\begin{gathered} 0.9172 \\ (1.2572) \\ \hline \end{gathered}$ | $\begin{gathered} 0.7430 \\ (1.2592) \\ \hline \end{gathered}$ |
| likelihood ratio test p-value | $\begin{gathered} \mathrm{Col}(1) \text { and (3) } \\ 0.0133 \end{gathered}$ | $\begin{gathered} \mathrm{Col}(2) \text { and (3) } \\ 0.0679 \end{gathered}$ |  |
| Number of observations | 688 | 688 | 688 |

[^16]Table 15: Under 65 years: collective vs unitary approach - information criteria. Heckman outcome equation: household share of net financial wealth allocated in risky assets. Partners characteristics by financial respondent.

| Under 65 years |  |  |
| :--- | :---: | :---: |
| share of risky assets | Collective model | Unitary model |
| demographics ${ }^{a}$ | $(1)$ | $(2)$ |
|  |  | $*$ |
| financial risk: hh | $0.0293^{* * *}$ |  |
|  | $(0.0074)$ |  |
| financial risk: respondent |  | $0.0225^{* * *}$ |
|  |  | $(0.0063)$ |
| Constant | 0.8925 | 0.7187 |
|  | $(1.2085)$ | $(1.2102)$ |
| AIC | $\mathbf{9 9 1}$ | $\mathbf{9 9 5}$ |
| BIC | $\mathbf{1 2 1 8}$ | $\mathbf{1 2 2 3}$ |
| Number of observations | 688 | 688 |

[^17]
[^0]:    *Acknowledgments: The author have benefited from the comments of Guglielmo Weber, Vincenzo Rettore, Chiara Dal Bianco, Francesca Parodi, Raffaele Miniaci, Mario Padula, and participants to the internal seminar series of the Department of Economics and Management of the University of Padova and the University of Modena and Reggio Emilia.
    The English Longitudinal Study of Ageing was developed by a team of researchers based at University College London, NatCen Social Research, the Institute for Fiscal Studies, the University of Manchester and the University of East Anglia. The data were collected by NatCen Social Research. The funding is currently provided by the National Institute on Aging (Ref: R01AG017644) and by a consortium of UK government departments: Department for Health and Social Care; Department for Transport; Department for Work and Pensions, which is coordinated by the National Institute for Health Research (NIHR, Ref: 198-1074). Funding has also been provided by the Economic and Social Research Council (ESRC).
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[^1]:    ${ }^{3}$ Mean-variance utility is equivalent to CARA utility, as I show in Section A. 1 Section A. 2 of the Appendix provides a second version of the model where household utility is a weighted sum of partners utility.

[^2]:    ${ }^{4}$ The assumption about homogeneous expectation allows to infer the implied degree of absolute risk aversion of the stockholders, i.e. only for those households whose share of wealth allocated in risky assets is observed.

[^3]:    5 Banks et al. (2021).The English Longitudinal Study of Ageing was developed by a team of researchers based at University College London, NatCen Social Research, the Institute for Fiscal Studies, the University of Manchester and the University of East Anglia. The data were collected by NatCen Social Research. The funding is currently provided by the National Institute on Aging in the US, and a consortium of UK government departments coordinated by the National Institute for Health Research. Funding has also been received by the Economic and Social Research Council.
    ${ }^{6}$ More information about HSE at http://healthsurvey.hscic.gov.uk

[^4]:    ${ }^{7}$ Financial risk tolerance question is: Thinking specifically about your finances, spending and savings, are you a person who is fully prepared to take risk, or do you try to avoid taking risks?

[^5]:    ${ }^{8}$ Financial patience question is: Thinking specifically about your finances, spending and savings, are you generally an impatient person, or someone who always shows great patience?
    ${ }^{9}$ Table 1 in the Appendix shows the sample selection procedure and the correspondent number of observations.
    ${ }^{10}$ ISA (Individual Saving Account) is a class of retail investment arrangement available to residents of the United Kingdom, with favorable tax condition. They offer four types of account: cash ISA, stocks \& shares ISA, innovative finance ISA (IFISA) and lifetime ISA.

[^6]:    ${ }^{11}$ Assume that there are two households with the same risk preferences and demographics, but different financial wealth. In other words, the households have the same optimal $\alpha$, but the total amount of money allocated in risky assets differ. The household with a higher financial wealth invests a higher amount in risky assets and obtains higher returns (in absolute terms), with a lower impact of fixed costs on its finances.

[^7]:    ${ }^{a}$ wealth is net total household wealth, housing is gross housing wealth (the value of owner-occupied primary housing before mortgage debt is subtracted), safe assets are money invested in "safe" assets such as bank accounts, savings accounts and cash ISAs, risky assets are money invested in "risky" assets such as shares, bonds, stocks and shares ISAs or life insurance ISAs, physical wealth represents alternative investments (second homes, farm or business property, works of art etc), debt is credit cards, overdrafts, other private debt but not mortgages.

[^8]:    ${ }^{12}$ Note that there is one household where the female reports negative earnings: in this case, I consider the male partner as the only income supplier with a share of household income equal to 1 . Therefore, there are 74 one-income couples where males are the only income source, while only 73 females with no income.

[^9]:    ${ }^{13}$ It is constructed with principal component, using the numerous questions of ELSA related to participants health conditions. For more information see Dal Bianco (2020).

[^10]:    ${ }^{14}$ I test the joint significance of the numeracy scores of the male and female partners in each model. The Wald tests reject the null hypothesis of joint non-significance at the $1 \%$ significance level.

[^11]:    ${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the health index of partners.

[^12]:    ${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the health index of partners.

[^13]:    ${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the health index of partners.

[^14]:    ${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the sickness index of partners.

[^15]:    ${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the sickness index of partners.

[^16]:    ${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the sickness index of partners.

[^17]:    ${ }^{a}$ Demographics include male age and age squared, dummy for large age difference between partners ( $>10$ years), dummies of income quartile, job market participation of partners, education of partners and the sickness index of partners.

